

## ISIPTA ‘09

Proceedings of the Sixth International Symposium on Imprecise Probability: Theories and Applications

Edited by
Thomas Augustin
Frank P. A. Coolen
Serafin Moral
Matthias C. M. Troffaes

## ISIPTA '09

Proceedings of the Sixth International Symposium on Imprecise Probability: Theories and Applications

Durham University, United Kingdom

14-18 July 2009

## Edited by

Thomas Augustin
Frank P. A. Coolen
Serafín Moral
Matthias C. M. Troffaes

Published by SIPTA
Society for Imprecise Probability: Theories and Applications
http://www.sipta.org
Cover, Copyright 2009 by Judith Aird.
Preface, Copyright 2009 by SIPTA.
Contributed papers, Copyright 2009 by their respective authors.
All rights reserved. The copyright on each of the papers published in these proceedings remains with the author(s). No part of these proceedings may be reprinted or reproduced or utilised in any form by any electronic, mechanical, or other means without permission in writing from the relevant author(s).

This book was typeset using $\mathrm{EAT}_{\mathrm{E}} \mathrm{X}$.

## Contents

Preface ..... vii
Organisation ..... xi
Iterated Random Selection as Intermediate Between Risk and Uncertainty Horacio Arló Costa, Jeffrey Helzner ..... 1
Closure of Independencies under Graphoid Properties: Some Experimental Re- sults
Marco Baioletti, Giuseppe Busanello, Barbara Vantaggi ..... 11
Category Selection for Multinomial Data
Rebecca Baker, Frank Coolen ..... 21
Aggregating Imprecise Probabilistic Knowledge
Alessio Benavoli, Alessandro Antonucci ..... 31
Tests of the Mean with Distributional Uncertainty: An Info-Gap Approach
Yakov Ben-Haim ..... 41
On General Conditional Random Quantities
Veronica Biazzo, Angelo Gilio, Giuseppe Sanfilippo ..... 51
Approximation of Coherent Lower Probabilities by 2-Monotone Measures
Andrew Bronevich, Thomas Augustin ..... 61
On the Use of a New Discrepancy Measure to Correct Incoherent Assessments and to Aggregate Conflicting Opinions Based on Imprecise Conditional Prob- abilities
Andrea Capotorti, Giuliana Regoli, Francesca Vattari ..... 71
A Generalization of Credal Networks
Marco Cattaneo ..... 79
A Tree Augmented Classifier Based on Extreme Imprecise Dirichlet Model
Giorgio Corani, Cassio Campos, Sun Yi ..... 89
Sets of Desirable Gambles and Credal Sets
Inés Couso, Serafín Moral ..... 99
Concentration Inequalities and Laws of Large Numbers under Epistemic Irrele- vance
Fabio Cozman ..... 109
Imprecise Markov Chains with an Absorbing State
Richard Crossman, Pauline Coolen-Schrijner, Damjan Skulj, Frank Coolen ..... 119
Credal Semantics of Bayesian Transformations
Fabio Cuzzolin ..... 129
Consistent Approximations of Belief Functions
Fabio Cuzzolin ..... 139
Epistemic Irrelevance in Credal Networks: The Case of Imprecise Markov Trees
Gert de Cooman, Filip Hermans, Alessandro Antonucci, Marco Zaffalon ..... 149
Exchangeability for Sets of Desirable Gambles Gert de Cooman, Erik Quaeghebeur ..... 159
Representing and Solving Factored Markov Decision Processes with Imprecise Probabilities
Karina Delgado, Leliane Barros, Fabio Cozman, Ricardo Shirota Filho ..... 169
The Role of Generalised p-Boxes in Imprecise Probability Models
Sébastien Destercke, Didier Dubois ..... 179
Boundary Linear Utility and Sensitivity of Decisions with Imprecise Utility Trade- Off Parameters
Malcolm Farrow, Michael Goldstein ..... 189
Multivariate Models and Confidence Intervals: A Local Random Set Approach Thomas Fetz ..... 199
A Minimum Distance Estimator in an Imprecise Probability Model: Computa- tional Aspects and Applications Robert Hable ..... 209
How Can We Get New Knowledge? Frank Hampel ..... 219
Dutch Books and Combinatorial Games
Peter Harremoës ..... 229
Characterizing Factuality in Normal Form Sequential Decision Making
Nathan Huntley, Matthias C. M. Troffaes ..... 239
Almost Probabilistic Assignments and Conditional Independence (a contribution to Dempster-Shafer theory of evidence) Radim Jiroušek ..... 249
On the Behavior of the Robust Bayesian Combination Operator and the Signifi- cance of Discounting Alexander Karlsson, Ronnie Johansson, Sten F. Andler ..... 259
Affinity and Continuity of Credal Set Operator Tomáš Kroupa ..... 269
Imprecise Probabilities from Imprecise Descriptions of Real Numbers Jonathan Lawry, Inés González-Rodríguez, Yongchuan Tang ..... 277
Reasoning with Imprecise Probabilistic Knowledge on Enzymes for Rapid Screen- ing of Potential Substrates or Inhibitor Structures Weiru Liu, Anbu Yue, David J. Timson ..... 287

## Organisation

## Steering Committee

Thomas Augustin, Germany
Frank P. A. Coolen, UK
Gert de Cooman, Belgium
Serafín Moral, Spain
Teddy Seidenfeld, US
Matthias C. M. Troffaes, UK

## Sponsors



## ELSEVIER

http://www.elsevier.com/

## EPSPC Engineering and Physical Sciences Research Council

http://www.epsrc.ac.uk/

[^0]
# Almost Bayesian Assignments and Conditional Independence (a contribution to Dempster-Shafer theory of evidence) 

Radim Jiroušek<br>Faculty of Management, University of Economics<br>Jindřichův Hradec<br>and<br>Institute of Information Theory and Automation<br>Academy of Sciences of the Czech Republic<br>radim@utia.cas.cz


#### Abstract

In the paper we introduce a family of almost Bayesian basic assignments, which slightly extends Bayesian basic assignments. This extension incorporates all the distributions that can be created from lowdimensional Bayesian basic assignments by application of the operator of composition, and simultaneously preserves the property of Bayesian basic assignments concerning the number of focal elements: it does not exceed cardinality of the frame of discernment. The other goal of the paper is to propagate a new way of definition of conditional independence relation in D-S theory. It follows ideas of P. P. Shenoy from [7], where the author defines the notion of conditional independence for valuation-based system based on his operation of "combination". Here we do the same but using the operator of "composition". The notion of independence we get in this way seems to meet better the general requirements on conditional independence relation for basic assignments.


Keywords. Dempster-Shafer theory of evidence, multidimensionality, operator of composition, conditional independence, semigraphoids.

## 1 Introduction

Regarding purely computational point of view, the greatest disadvantage of Dempster-Shafer theory of evidence (D-S) is that in contrast to probabilistic or possibilistic models, which can be described by the respective density functions (i.e. point functions), D-S models must be described by set functions. It means that while the number of necessary parameters of probabilistic or possibilistic models grows exponentially with the number of dimensions, for D-S models one needs a superexponential number of parameters.

It is known from theory of Bayesian networks (or graphical Markov models, in general) that the number of parameters can be drastically decreased by uti-
lization of properties of conditional independence relations valid for the modelled situation. This was among the reasons why we designed an alternative approach for multidimensional probability distribution representation based on so called operator of composition [2]. The basic idea of these models is very simple: multidimensional models are assembled (composed) from a system of low-dimensional distributions by the operator of composition (in a specified order). Later on, Vejnarová introduced an analogous operator also for composition of possibility distributions and showed it manifested similar properties as its probabilistic counterpart [10, 11]. Recently we designed the operator of composition also for basic assignments in D-S theory of evidence [5] and proved that it met all the required properties necessary for multidimensional models representation $[3,4]$.

However, it is not the goal of this paper to publicize advantageous properties of the operator of composition for basic assignments. The goal of this contribution is twofold. The first one is to show that there exists a family of basic assignments, for specification of which one does not need more parameters than for probabilistic models and yet it enables modelling some type of ignorance (Section 4). The other goal is to show that if the conditional independence for basic assignments is defined with the help of the operator of composition (which was already done in [3]) one can prove semigraphoid axioms from a small number of operator's basic properties. This is done in Section 5 .

## 2 Basic notion

## Set notation

In the whole paper we shall deal with a finite number of variables $X_{1}, X_{2}, \ldots, X_{n}$ each of which is specified by a finite set $\mathbf{X}_{i}$ of its values. So, we will consider multidimensional space of discernment

$$
\mathbf{X}_{N}=\mathbf{X}_{1} \times \mathbf{X}_{2} \times \ldots \times \mathbf{X}_{n}
$$

and its subspaces. For $K \subset N=\{1,2, \ldots, n\}, \mathbf{X}_{K}$ denotes a Cartesian product of those $\mathbf{X}_{i}$, for which $i \in K$ :

$$
\mathbf{X}_{K}=\mathbf{X}_{i \in K} \mathbf{X}_{i}
$$

A projection of $x=\left(x_{1}, x_{2}, \ldots, x_{n}\right) \in \mathbf{X}_{N}$ into $\mathbf{X}_{K}$ will be denoted $x^{\downarrow K}$, i.e. for $K=\left\{i_{1}, i_{2}, \ldots, i_{\ell}\right\}$

$$
x^{\downarrow K}=\left(x_{i_{1}}, x_{i_{2}}, \ldots, x_{i_{\ell}}\right) \in \mathbf{X}_{K}
$$

Analogously, for $K \subset L \subseteq N$ and $A \subset \mathbf{X}_{L}, A^{\downarrow K}$ will denote a projection of $A$ into $\mathbf{X}_{K}$ :

$$
A^{\downarrow K}=\left\{y \in \mathbf{X}_{K}: \exists x \in A \quad\left(y=x^{\downarrow K}\right)\right\} .
$$

Let us remark that we do not exclude situations when $K=\emptyset$. In this case $A^{\downarrow \emptyset}=\emptyset$.

Set $A \subseteq \mathbf{X}_{K}$ is said to be a point-cylinder if it can be expressed as a Cartesian product of a singleton and a subspace $\mathbf{X}_{L}$. More precisely: a point-cylinder is a set $A \subseteq \mathbf{X}_{K}$ for which there exists an index set (possibly empty) $L \subseteq K$ such that $\left|C^{\downarrow L}\right| \leq 1$ and

$$
C=C^{\downarrow L} \times \mathbf{X}_{K \backslash L}
$$

Let us stress that if $L=\emptyset$ then $C=\mathbf{X}_{K}$ (it is the only situation when $\left|C^{\downarrow L}\right|<1$ ), and when $L=K$ then $|C|=1$.

In addition to the projection, in this text we will need also the opposite operation which will be called a join. By a join of two sets $A \subseteq \mathbf{X}_{K}$ and $B \subseteq \mathbf{X}_{L}$ we will understand a set

$$
A \otimes B=\left\{x \in \mathbf{X}_{K \cup L}: x^{\downarrow K} \in A \quad \& \quad x^{\downarrow L} \in B\right\}
$$

Notice that if $K$ and $L$ are disjoint then the join of the corresponding sets is just their Cartesian product

$$
A \otimes B=A \times B
$$

For $K=L, A \otimes B=A \cap B$. If $K \cap L \neq \emptyset$ and $A^{\downarrow K \cap L} \cap B^{\downarrow K \cap L}=\emptyset$ then also $A \otimes B=\emptyset$.

In one of the following proofs we will need the following (rather technical) property of set joins.

Lemma 1. Let $K_{1} \cap K_{2} \subseteq L \subseteq K_{2} \subseteq N$. Then for any $C \subseteq \mathbf{X}_{K_{1} \cup K_{2}}$ the following condition (a) holds if and only if both conditions (b) and (c) hold true.
(a) $C=C^{\downarrow K_{1}} \otimes C^{\downarrow K_{2}}$;
(b) $C^{\downarrow K_{1} \cup L}=C^{\downarrow K_{1}} \otimes C^{\downarrow L}$;
(c) $C=C^{\downarrow K_{1} \cup L} \otimes C^{\downarrow K_{2}}$.

Proof. Let us prove the assertion in three steps. First, however, let us realize that

$$
x \in C \Longrightarrow\left(x^{\downarrow K_{1}} \in C^{\downarrow K_{1}} \& x^{\downarrow K_{2}} \in C^{\downarrow K_{2}}\right)
$$

and therefore $C=C^{\downarrow K_{1}} \otimes C^{\downarrow K_{2}}$ is equivalent to

$$
\begin{aligned}
\forall x \in & \mathbf{X}_{K_{1} \cup K_{2}} \\
& \left(x^{\downarrow K_{1}} \in C^{\downarrow K_{1}} \& x^{\downarrow K_{2}} \in C^{\downarrow K_{2}} \Longrightarrow x \in C\right) .
\end{aligned}
$$

(a) $\Longrightarrow(b)$.

Consider $x \in \mathbf{X}_{K_{1} \cup L}$, such that $x^{\downarrow K_{1}} \in C^{\downarrow K_{1}}$ and $x^{\downarrow L} \in C^{\downarrow L}$. Since $x^{\downarrow L} \in C^{\downarrow L}$ there must exists (at least one) $y \in C^{\downarrow K_{2}}$, for which $y^{\downarrow L}=x^{\downarrow L}$. Now construct $z \in \mathbf{X}_{K_{1} \cup K_{2}}$ for which $z^{\downarrow K_{1}}=x^{\downarrow K_{1}}$ and $z^{\downarrow K_{2}}=y$ (it is possible because $y^{\downarrow L}=x^{\downarrow L}$ ). From this construction we see that $z^{\downarrow K_{1} \cup L}=x$. Therefore $z^{\downarrow K_{1}}=x^{\downarrow K_{1}} \in C^{\downarrow K_{1}}$ and $z^{\downarrow K_{2}}=y \in C^{\downarrow K_{2}}$ form which, because we assume that (a) holds, we get that $z \in C$, and therefore also $x=z^{\downarrow K_{1} \cup L} \in C^{\downarrow K_{1} \cup L}$.
(a) $\Longrightarrow$ (c).

Consider now $x \in \mathbf{X}_{K_{1} \cup K_{2}}$, for which its projections $x^{\downarrow K_{1} \cup L} \in C^{\downarrow K_{1} \cup L}$ and $x^{\downarrow K_{2}} \in C^{\downarrow K_{2}}$. From $x^{\downarrow K_{1} \cup L} \in$ $C^{\downarrow K_{1} \cup L}$ we immediately get that $x^{\downarrow K_{1}} \in C^{\downarrow K_{1}}$, which in combination with $x^{\downarrow K_{2}} \in C^{\downarrow K_{2}}$ (due to the assumption (a)) yields that $x \in C$.
(b) \& (c) $\Longrightarrow(\mathrm{a})$.

Consider $x \in \mathbf{X}_{K_{1} \cup K_{2}}$ such that $x^{\downarrow K_{1}} \in C^{\downarrow K_{1}}$ and $x^{\downarrow K_{2}} \in C^{\downarrow K_{2}}$. From the last property one gets also $x^{\downarrow L} \in C^{\downarrow L}$, which, in combination with $x^{\downarrow K_{1}} \in C^{\downarrow K_{1}}$ gives, because (b) holds true, that $x^{\downarrow K_{1} \cup L} \in C^{\downarrow K_{1} \cup L}$. And the last property in combination with $x^{\downarrow K_{2}} \in$ $C^{\downarrow K_{2}}$ yields the required $x \in C$.

## Assignment notation

The role of a probability distribution from a probability theory is in Dempster-Shafer theory played by any of the set functions: belief function, plausibility function or basic (probability or belief) assignment. Knowing one of them, one can deduce the two remaining. In this paper we shall use exclusively basic assignments.

A basic assignment $m$ on $\mathbf{X}_{K}(K \subseteq N)$ is a function

$$
m: \mathcal{P}\left(\mathbf{X}_{K}\right) \longrightarrow[0,1]
$$

for which

$$
\sum_{\emptyset \neq A \subseteq \mathbf{x}_{N}} m(A)=1
$$

For the sake of this paper it is reasonable to consider only normalized basic assignments, for which $m(\emptyset)$ equals always 0 . If $m(A)>0$, then $A$ is said to be a focal element of $m$.

Having a basic assignment $m$ on $\mathbf{X}_{K}$ one can consider its marginal assignment on $\mathbf{X}_{L}$ (for $L \subseteq K$ ), which is defined (for each $\emptyset \neq B \subseteq \mathbf{X}_{L}$ ):

$$
m^{\downarrow L}(B)=\sum_{A \subseteq \mathbf{X}_{K}: A^{\downarrow L}=B} m(A)
$$

Basic assignment $m$ is said to be Bayesian if all its focal elements are singletons, i.e.

$$
m(A)>0 \quad \Longrightarrow \quad|A|=1
$$

In this case, namely, both the other two functions, belief Bel and plausibility Pl which are defined by the following formulas (for all $A \subseteq \mathbf{X}_{K}$ )

$$
\begin{aligned}
& \operatorname{Bel}(A)=\sum_{B \subseteq A} m(A) \\
& P l(A)=1-\operatorname{Bel}(\bar{A})
\end{aligned}
$$

are normalized additive functions, and therefore probability distributions.
Another special case is represented by simple basic assignments. Basic assignments $m$ on $\mathbf{X}_{K}$ is called simple if there exists $A\left(\emptyset \neq A \subset \mathbf{X}_{K}\right)$ and a positive number $a$ such that $m(A)=a$ and $m\left(\mathbf{X}_{K}\right)=1-a$.

## 3 Operator of composition

Originally, the operator of composition was designed in probability theory as a tool enabling creation of multidimensional probability distributions - multidimensional models - by successive composition of lowdimensional distributions. The basic idea of this operator was simple. It generalized the fact that one can construct a 3-dimensional probability distribution $P(X, Y, Z)$ from two 2-dimensional ones $Q(X, Y)$ and $R(Y, Z)$ just by assigning

$$
P(X, Y, Z)=Q(X, Y) \cdot R(Z \mid Y)
$$

In this case $P$ reflects all the information contained in $Q$, because evidently $P(X, Y)=Q(X, Y)$, and some of the information contained in $R(P(Z \mid Y)=$ $R(Z \mid Y)$ ). Moreover, $P$ does not contain any additional information, because for this probability distribution variables $X$ and $Z$ are conditionally independent given variable $Y$.

Introduction of the probabilistic operator of composition opened a study of a new area called compositional models, which was an alternative to Bayesian networks, or to Graphical Markov models in general. Though it appeared that Bayesian networks and compositional models described exactly the same class of probability distributions, study of a new type of
models appeared useful. First of all it offered new points of view to multidimensional probability distribution representation. In addition to this, compositional models were in some situations more advantageous from the computational point of view (some of the marginal distributions, computation of which may be algorithmically rather expensive, were in a compositional model expressed explicitly).
Later, the operator of composition was designed and studied in possibility theory by Vejnarová [10]. Being inspired by Didier Dubois, we introduced the operator of composition also for basic assignments [5]; this definition is presented below. In that paper we also showed that if the operator of composition is applied to Bayesian basic assignments it usually yields the Bayesian basic assignment, which corresponds to the probability distribution, which is constructed by the probabilistic operator of composition from the respective probability distributions. The only exception from this situation occurs when composing basic assignments corresponding to probability distributions, for which their probabilistic composition is not defined. In such a case, result of composition of such Bayesian basic assignments is not Bayesian. In the next section we will reveal the main characteristics of such basic assignments.

Definition 1. For two arbitrary basic assignments $m_{1}$ on $\mathbf{X}_{K}$ and $m_{2}$ on $\mathbf{X}_{L}(K \neq \emptyset \neq L)$ a composition $m_{1} \triangleright m_{2}$ is defined for each $C \subseteq \mathbf{X}_{K \cup L}$ by one of the following expressions:
[a] if $m_{2}^{\downarrow K \cap L}\left(C^{\downarrow K \cap L}\right)>0$ and $C=C^{\downarrow K} \otimes C^{\downarrow L}$ then

$$
\left(m_{1} \triangleright m_{2}\right)(C)=\frac{m_{1}\left(C^{\downarrow K}\right) \cdot m_{2}\left(C^{\downarrow L}\right)}{m_{2}^{\downarrow K \cap L}\left(C^{\downarrow K \cap L}\right)} ;
$$

[b] if $m_{2}^{\downarrow K \cap L}\left(C^{\downarrow K \cap L}\right)=0$ and $C=C^{\downarrow K} \times \mathbf{X}_{L \backslash K}$ then

$$
\left(m_{1} \triangleright m_{2}\right)(C)=m_{1}\left(C^{\downarrow K}\right) ;
$$

[c] in all other cases $\left(m_{1} \triangleright m_{2}\right)(C)=0$.

Before illustrating the operator of composition on a simple example, let us remark that three expressions in Definition 1 correspond to three situations, which occur when one wants to define a basic assignments possessing those properties we highlighted when speaking about the probability distribution $P(X, Y, Z)=Q(X, Y) \cdot R(Z \mid Y)$. Point [a], in a way, directly corresponds to this well-known probabilistic formula. It disseminates the mass $m_{1}\left(C^{\downarrow K}\right)$ into the respective subsets $C \subseteq \mathbf{X}_{K \cup L}$. The information describing the way how this mass is disseminated is taken over from $m_{2}$. Point [b] is applicable when

Table 1: 1-dimensional basic assignments $m_{1}$ and $m_{2}$.

| $A \subseteq \mathbf{X}_{1}$ | $m_{1}(A)$ |
| :---: | :---: |
| $\{a\}$ | 0.5 |
| $\{\bar{a}\}$ | 0.1 |
| $\{a, \bar{a}\}$ | 0.4 |

$m_{2}^{\downarrow K \cap L}\left(C^{\downarrow K \cap L}\right)=0$ and therefore $m_{2}$ does not determine the way how to disseminate the respective mass. Therefore the whole mass $m_{1}\left(C^{\downarrow K}\right)$ is assigned to the least specific set: $C=C^{\downarrow K} \times \mathbf{X}_{L \backslash K}$ (expressing in this way maximal ignorance). Eventually, point [c] guarantees that no additional information is added to the resulting basic assignment $m_{1} \triangleright m_{2}$. It assigns zero mass to all those subsets of $\mathbf{X}_{K \cup L}$, whose positive values would violate the notion of the required conditional independence (see e.g. [1]).

Example 1. Consider two 1-dimensional basic assignments ${ }^{1} m_{1}, m_{2}$ from Table 1, which are defined on $\mathbf{X}_{1}=\{a, \bar{a}\}$ and $\mathbf{X}_{2}=\{b, \bar{b}\}$, respectively.
Their composition $m_{1} \triangleright m_{2}$ is in Table 2. Notice, that this composed basic assignment has only 6 focal elements, which means that for the remaining $\left(2^{4}-1\right)-6=9$ subsets of $\mathbf{X}_{1} \times \mathbf{X}_{2}$, values of $m_{1} \triangleright m_{2}$ equal 0 . It is the case of two groups of subsets. As for three subsets

$$
\begin{aligned}
& \{a b, a \bar{b}\}=\{a\} \otimes \mathbf{X}_{2}, \\
& \{\bar{a} b, \bar{a} \bar{b}\}=\{\bar{a}\} \otimes \mathbf{X}_{2}, \\
& \{a b, a \bar{b}, \bar{a} b, \bar{a} \bar{b}\}=\mathbf{X}_{1} \otimes \mathbf{X}_{2},
\end{aligned}
$$

their values of $m_{1} \triangleright m_{2}$ are assigned by point [a] of Definition 1 and equal 0 because $m_{2}(\{b, \bar{b}\})=0$. On the other hand side, to the remaining six subsets

$$
\begin{aligned}
& \{a b, \bar{a} \bar{b}\}, \\
& \{a \bar{b}, \bar{a} b\}, \\
& \{a b, a \bar{b}, \bar{a} b\}, \\
& \{a b, a \bar{b}, \bar{a} \bar{b}\}, \\
& \{a b, \bar{a} b, \bar{a} \bar{b}\}, \\
& \{a \bar{b}, \bar{a} b, \bar{a} \bar{b}\},
\end{aligned}
$$

values of $m_{1} \triangleright m_{2}$ are assigned by point [c] of Definition 1, because for these subsets it does not hold that $C=C^{\downarrow\{1\}} \otimes C^{\downarrow\{2\}}$. Assigning a positive value to any of these subsets we would, in a way, introduce a dependence of variables $X_{1}$ and $X_{2}$.

[^1]Table 2: Composed basic assignment $m_{1} \triangleright m_{2}$.

| $C \subseteq \mathbf{X}_{1} \times \mathbf{X}_{2}$ | $\left(m_{1} \triangleright m_{2}\right)(C)$ |
| :---: | :---: |
| $\{a b\}$ | 0.25 |
| $\{a \bar{b}\}$ | 0.25 |
| $\{\bar{a} b\}$ | 0.05 |
| $\{\bar{a} \bar{b}\}$ | 0.05 |
| $\{a b, \bar{a} b\}$ | 0.20 |
| $\{a \bar{b}, \bar{a} \bar{b}\}$ | 0.20 |

Let us present the most important properties of the operator of composition for basic assignments.

Lemma 2. Let $K, L \subseteq N$. For arbitrary basic assignments $m_{1}, m_{2}$ defined on $\mathbf{X}_{K}, \mathbf{X}_{L}$, respectively
(i) $m_{1} \triangleright m_{2}$ is a basic assignment on $\mathbf{X}_{K \cup L}$;
(ii) $\left(m_{1} \triangleright m_{2}\right)^{\downarrow K}=m_{1}$;
(iii) $m_{1} \triangleright m_{2}=m_{2} \triangleright m_{1} \quad \Longleftrightarrow \quad m_{1}^{\downarrow K \cap L}=m_{2}^{\downarrow K \cap L}$;
(iv) $L \supseteq M \supseteq(K \cap L)$

$$
\Longrightarrow \quad m_{1} \triangleright m_{2}=\left(m_{1} \triangleright m_{2}^{\downarrow M}\right) \triangleright m_{2} \text {; }
$$

Proof. The first three properties were proved in [5]: properties (i)-(iii) are properties (i)-(iii) of Lemma 1. Thus, what has remained to be proved is just property (iv).

So, our goal is to show that for basic assignments $m_{1}, m_{2}$ and for any $M$ such that $L \supseteq M \supseteq K \cap L$

$$
\left(m_{1} \triangleright m_{2}\right)(C)=\left(\left(m_{1} \triangleright m_{2}^{\downarrow M}\right) \triangleright m_{2}\right)(C) .
$$

holds true for any $C \subseteq \mathbf{X}_{K \cup L}$.
The proof will be performed in three steps corresponding to cases [a], [b], [c] of Definition 1.
Ad [a]. Assume that $C=C^{\downarrow K} \otimes C^{\downarrow L}$ and $m_{2}^{\downarrow K \cap L}\left(C^{\downarrow K \cap L}\right)>0$. From this we get from Lemma 1 that also $C^{\downarrow K \cup M}=C^{\downarrow K} \otimes C^{\downarrow M}$, and therefore (since $K \cap L=K \cap M)$

$$
\left(m_{1} \triangleright m_{2}^{\downarrow M}\right)\left(C^{\downarrow K \cup M}\right)=\frac{m_{1}\left(C^{\downarrow K}\right) \cdot m_{2}^{\downarrow M}\left(C^{\downarrow M}\right)}{m_{2}^{\downarrow K \cap L}\left(C^{\downarrow K \cap L}\right)} .
$$

In the rest of this step we have to distinguish two situations depending whether $m_{2}^{\downarrow M}\left(C^{\downarrow M}\right)$ equals 0 or not.
If $m_{2}^{\downarrow M}\left(C^{\downarrow M}\right)>0$ (realize that in this case also
$\left.m_{2}^{\downarrow K \cap L}\left(C^{\downarrow K \cap L}\right)>0\right)$ then

$$
\begin{aligned}
& \left(\left(m_{1} \triangleright m_{2}^{\downarrow M}\right) \triangleright m_{2}\right)(C) \\
& \quad=\frac{\left(m_{1} \triangleright m_{2}^{\downarrow M}\right)\left(C^{\downarrow K \cup M}\right) \cdot m_{2}\left(C^{\downarrow L}\right)}{m_{2}^{\downarrow M}\left(C^{\downarrow M}\right)} \\
& \quad=\frac{\frac{m_{1}\left(C^{\downarrow K}\right) \cdot m_{2}^{\downarrow M}\left(C^{\downarrow M}\right)}{m_{2}^{\downarrow K \cap L}\left(C^{\downarrow K \cap L}\right)} \cdot m_{2}\left(C^{\downarrow L}\right)}{m_{2}^{\downarrow M}\left(C^{\downarrow M}\right)} \\
& \quad=\frac{m_{1}\left(C^{\downarrow K}\right) \cdot m_{2}\left(C^{\downarrow L}\right)}{m_{2}^{\downarrow K \cap L}\left(C^{\downarrow K \cap L}\right)}=\left(m_{1} \triangleright m_{2}\right)(C) .
\end{aligned}
$$

If $m_{2}^{\downarrow M}\left(C^{\downarrow M}\right)=0$ then, according to Definition 1, either

$$
\left(\left(m_{1} \triangleright m_{2}^{\downarrow M}\right) \triangleright m_{2}\right)(C)=\left(m_{1} \triangleright m_{2}^{\downarrow M}\right)\left(C^{\downarrow K \cup M}\right),
$$

in case that $C=C^{\downarrow K \cup M} \otimes \mathbf{X}_{L \backslash M}$, or

$$
\left(\left(m_{1} \triangleright m_{2}^{\downarrow M}\right) \triangleright m_{2}\right)(C)=0,
$$

in opposite case. However, in this case also

$$
\left(m_{1} \triangleright m_{2}^{\downarrow M}\right)\left(C^{\downarrow K \cup M}\right)=\frac{m_{1}\left(C^{\downarrow K}\right) \cdot m_{2}^{\downarrow M}\left(C^{\downarrow M}\right)}{m_{2}^{\downarrow K \cap L}\left(C^{\downarrow K \cap L}\right)}=0,
$$

and therefore $\left(\left(m_{1} \triangleright m_{2}^{\perp M}\right) \triangleright m_{2}\right)(C)=0$ regardless of the form of $C^{\downarrow L \backslash M}$ (i.e. for both situations: $C^{\downarrow L \backslash M}=\mathbf{X}_{L \backslash M}$ and $\left.C^{\downarrow L \backslash M} \neq \mathbf{X}_{L \backslash M}\right)$. Taking into consideration the fact that in the considered situation (i.e. $m_{2}^{\downarrow M}\left(C^{\downarrow M}\right)=0$ ) also $m_{2}\left(C^{\downarrow L}\right)=0$, and therefore also

$$
\left(m_{1} \triangleright m_{2}\right)(C)=\frac{m_{1}\left(C^{\downarrow K}\right) \cdot m_{2}\left(C^{\downarrow L}\right)}{m_{2}^{\downarrow K \cap L}(C \downarrow K \cap L)}=0,
$$

we have finished the first step of the proof.
$\operatorname{Ad}[\mathrm{b}]$. Now we assume that $C=C^{\downarrow K} \otimes \mathbf{X}_{L \backslash K}$, and that $m_{2}^{\downarrow K \cap L}\left(C^{\downarrow K \cap L}\right)=0$. In this case, naturally, also $m_{2}^{\downarrow M}\left(C^{\downarrow M}\right)=0$ and $C=C^{\downarrow K} \otimes \mathbf{X}_{M \backslash K} \otimes \mathbf{X}_{L \backslash M}$. Therefore, according to case [b] of Definition 1,

$$
\left(m_{1} \triangleright m_{2}^{\downarrow M}\right)\left(C^{\downarrow K \cup M}\right)=m_{1}\left(C^{\downarrow K}\right),
$$

and because of the same reasons also

$$
\begin{aligned}
\left(\left(m_{1} \triangleright m_{2}^{\downarrow M}\right) \triangleright m_{2}\right)(C) & =\left(m_{1} \triangleright m_{2}^{\downarrow M}\right)\left(C^{\downarrow K \cup M}\right) \\
& =m_{1}\left(C^{\downarrow K}\right) .
\end{aligned}
$$

In this case also $\left(m_{1} \triangleright m_{2}\right)(C)=m_{1}\left(C^{\downarrow K}\right)$, and we have finished the second step of the proof.
$\operatorname{Ad}[c]$. The last step is trivial. In this case, as the reader can immediately see, both $\left(\left(m_{1} \triangleright m_{2}^{\downarrow M}\right) \triangleright\right.$ $\left.m_{2}\right)(C)$ and $\left(m_{1} \triangleright m_{2}\right)(C)$ equal 0 and therefore they equal to each other.

Table 3: 2-dimensional basic assignments $m_{3}$ and $m_{4}$.

| $A \subseteq \mathbf{X}_{\{1,2\}}$ | $m_{3}(A)$ |  | $B \subseteq \mathbf{X}_{\{2,3\}}$ | $m_{4}(B)$ |
| :---: | :---: | :---: | :---: | :---: |
| $\{a \bar{b}\}$ | 0.5 |  | $\{b c\}$ | 0.5 |
| $\{\bar{a} b\}$ | 0.1 |  | $\{\bar{b} \bar{c}\}$ | 0.2 |
| $\{a b, \bar{a} b\}$ | 0.4 |  | $\{b \bar{c}, \bar{b} c\}$ | 0.3 |

Table 4: Basic assignments $m_{3} \triangleright m_{4}$ and $m_{4} \triangleright m_{3}$.

| $C \subseteq \mathbf{X}_{1} \times \mathbf{X}_{2} \times \mathbf{X}_{3}$ | $m_{3} \triangleright m_{4}$ | $m_{4} \triangleright m_{3}$ |
| :---: | :---: | :---: |
| $\{a \bar{b} \bar{c}\}$ | 0.5 | 0.2 |
| $\{\bar{a} b c\}$ | 0.1 | 0.1 |
| $\{a b c, \bar{a} b c\}$ | 0.4 | 0.4 |
| $\{a b \bar{c}, a \bar{b} c, \bar{a} b \bar{c}, \bar{a} \bar{b} c\}$ |  | 0.3 |

Example 2. Property (iii) of the previous lemma says that for consistent basic assignments the operator of composition is commutative. Since any couple of basic assignments defined on non-overlapping frames of discernment are consistent (because $m^{\downarrow \emptyset}=$ 1 ), for basic assignments $m_{1}$ and $m_{2}$ from Table 1 $m_{1} \triangleright m_{2}=m_{2} \triangleright m_{1}$. Therefore, if we want to illustrate non-commutativity of this operator we have to consider overlapping frames of discernment ${ }^{2}$.
Consider basic assignments $m_{3}, m_{4}$ from Table 3. The reader can easily see that when computing $m_{3} \triangleright m_{4}$, all the focal elements are computed according to case [a] of Definition 1. There are only three sets $C \subseteq$ $\mathbf{X}_{\{1,2,3\}}$, for which $C=C^{\downarrow\{1,2\}} \otimes C^{\downarrow\{2,3\}}$, and for which both $m_{3}\left(C^{\downarrow\{1,2\}}\right)$ and $m_{3}\left(C^{\downarrow\{2,3\}}\right)$ are positive, namely

$$
\begin{aligned}
& \{a \bar{b} \bar{c}\}=\{a \bar{b}\} \otimes\{\bar{b} \bar{c}\} \\
& \{\bar{a} b c\}=\{\bar{a} b\} \otimes\{b c\} \\
& \{a b c, \bar{a} b c\}=\{a b, \bar{a} b\} \otimes\{b c\}
\end{aligned}
$$

On the other hand, when computing $m_{4} \triangleright m_{3}$ there appears set $C=\{b \bar{c}, \bar{b} c\} \times \mathbf{X}_{1}$, for which $m_{3}\left(C^{\downarrow\{1,2\}}\right)=0$ and therefore value $\left(m_{4} \triangleright m_{3}\right)(C)$ is assigned by point [b] of Definition 1. The resulting basic assignment $m_{4} \triangleright m_{3}$ is also recorded in Table 4.

Remark: In previous papers [5, 4] we showed a number of other properties of the operator of composition

[^2]for basic assignments, especially those useful for construction of multidimensional models. The four properties included in the previous lemma are those, which are sufficient to prove that conditional independence, if introduced with the help of the operator of composition (as done in Section 5), meets the semigraphoid axioms. In a way it is surprising that such a small group of elementary properties is sufficient. In connection with this fact a question arises whether the presented four properties are independent, whether some of them cannot be deduced from the remaining four.

Remark: Let us briefly answer a frequent question what is the relation of the introduced operator of composition and the famous Dempster's rule of combination ${ }^{3}$. Let us stress that the main difference emerges from the different purposes the operators where designed for. While Dempster's rule of combination was designed to have a tool enabling fusion of two basic assignments (the goal is to get a better information about the object than those contained in any of the original basic assignments), the operator of composition combines different descriptions of the object to comprehend all the information contained in original sources. This process corresponds to knowledge integration rather than knowledge fusion.
From the formal point of view this difference is reflected in property (ii) of Lemma 2, which holds for Dempster's rule of combination only in very specific (degenerated) situations. By the way, this difference is also the main reason why we consider the attempts to define a notion of conditional independence with the help of Dempster's rule of combination to be misleading.

## 4 Almost Bayesian basic assignments

One of the reasons (and from our point of view perhaps the most important) why D-S theory of evidence was designed and why it is in the center of attention of many researchers is the fact that probability theory has difficulties with representing some types of uncertainty; here we have in mind especially ignorance. For example, probability theory can hardly distinguish situation when an integer from $\{1,2, \ldots, 6\}$ is determined by tossing a fair die, and when it is selected by a totally unknown mechanism (well, the second situation can be described by the set of all possible distributions, however it is rather inconvenient). On the other hand, D-S theory yields very complex models and the corresponding computational procedures are of extremely high algorithmic complexity. Now,

[^3]we are about to specify a small family of basic assignments extending the set of Bayesian assignments but keeping the computational complexity on the level of probabilistic models. However, we have to admit that this new family, elements of which will be called almost Bayesian basic assignments, is very restrictive.

Definition 2. Basic assignment $m$ on $\mathbf{X}_{K}$ is called cylindrical if all its focal elements are point-cylinders.

Theorem 1. Let $K, L \subseteq N$ and $m_{1}, m_{2}$ be basic assignments defined on $\mathbf{X}_{K}$ and $\mathbf{X}_{L}$, respectively. If $m_{1}, m_{2}$ are cylindrical then $m_{1} \triangleright m_{2}$ is also cylindrical.

Proof. To prove this assertion we have to realize that a projection $A^{\downarrow K}$ of a point-cylinder $A$ is a pointcylinder. Moreover, join $A \otimes B$ of two point-cylinders $A$ and $B$ is again a point-cylinder (recall that $\emptyset$ is a point-cylinder).

Values of focal elements of basic assignment are computed according to either point [a] or point [b] of Definition 1. In case [a], a positive value can be assigned only if $C=C^{\downarrow K} \otimes C^{\downarrow L}$ and both $C^{\downarrow K}$ and $C^{\downarrow L}$ are point-cylinders. Case [b] is applied only when $C=C^{\downarrow K} \times \mathbf{X}_{L \backslash K}$. So in both cases positive value can be assigned only to point-cylinders.

Definition 3. Basic assignment $m$ on $\mathbf{X}_{K}$ is sparse if all its focal elements are pairwise disjoint.

Theorem 2. Let $K, L \subseteq N$ and $m_{1}, m_{2}$ be basic assignments defined on $\mathbf{X}_{K}$ and $\mathbf{X}_{L}$, respectively. If $m_{1}, m_{2}$ are sparse then $m_{1} \triangleright m_{2}$ is also sparse.

Proof. Consider two non-disjoint focal elements $C_{1}, C_{2}$ of $m_{1} \triangleright m_{2}:\left(m_{1} \triangleright m_{2}\right)\left(C_{1}\right)>0$ and $\left(m_{1} \triangleright\right.$ $\left.m_{2}\right)\left(C_{2}\right)>0$. Since $m_{1}$ is marginal of $m_{1} \triangleright m_{2}$, it is obvious that $C_{1}^{\downarrow K}$ and $C_{2}^{\downarrow K}$ are focal elements of $m_{1}$. Since we assume that $C_{1}$ and $C_{2}$ are non-disjoint the same must hold also for their projections

$$
C_{1}^{\downarrow K} \cap C_{2}^{\downarrow K} \neq \emptyset
$$

and therefore, because of our assumption that $m_{1}$ is sparse, $C_{1}^{\downarrow K}=C_{2}^{\downarrow K}$.
What are the focal elements $C$ of $m_{1} \triangleright m_{2}$, for which $C^{\downarrow K}=C_{1}^{\downarrow K}$ ? The answer to this question is offered by Definition 1 (realize that since we are considering focal elements $C$, values $\left(m_{1} \triangleright m_{2}\right)(C)$ are defined by expressions in points [a] or [b]).
If $m_{2}^{\downarrow K \cap L}\left(C^{\downarrow K \cap L}\right)>0$ then the considered focal elements can be expressed in the form

$$
C=C^{\downarrow K} \otimes C^{\downarrow L}=C_{1}^{\downarrow K} \otimes D
$$

Table 5: Sparse basic assignment $m$ on $\mathbf{X}_{\{1,2\}}$.

| $A \subseteq \mathbf{X}_{1} \times \mathbf{X}_{2}$ | $m(A)$ |
| :---: | :---: |
| $\{a b\}$ | 0.2 |
| $\{\bar{a} b\}$ | 0.3 |
| $\{a \bar{b}, \bar{a} \bar{b}\}$ | 0.5 |

Table 6: Marginal basic assignments $m^{\downarrow\{1\}}, m^{\downarrow\{2\}}$.

| $A \subseteq \mathbf{X}_{1}$ | $m^{\downarrow\{1\}}(A)$ |
| :---: | :---: |
| $\{a\}$ | 0.2 |
| $\{\bar{a}\}$ | 0.3 |
| $\{a, \bar{a}\}$ | 0.5 |


| $B \subseteq \mathbf{X}_{2}$ | $m^{\downarrow\{2\}}(B)$ |
| :---: | :---: |
| $\{b\}$ | 0.5 |
| $\{\bar{b}\}$ | 0.5 |

where $D \subseteq \mathbf{X}_{L}$ is a focal element of $m_{2}$ and $D^{\downarrow K \cap L}=$ $C_{1}^{\downarrow K \cap L}$. From this one can immediately see that $C_{1}=C_{1}^{\downarrow K} \otimes C_{1}^{\downarrow L}$ and $C_{2}=C_{1}^{\downarrow K} \otimes C_{2}^{\downarrow L}$ are disjoint if and only if also focal elements $C_{1}^{\downarrow L}$ and $C_{2}^{\downarrow L}$ of $m_{2}$ are disjoint. In our case, because $m_{2}$ is sparse, and because we assume that $C_{1} \cap C_{2} \neq \emptyset$, it means that $C_{1}^{\downarrow L}=C_{2}^{\downarrow L}$, and therefore also $C_{1}=C_{2}$.
In case that $m_{2}^{\downarrow K \cap L}\left(C^{\downarrow K \cap L}\right)=0$ then the situation is even simpler because in this case there can be only one focal element $C=C_{1}^{\downarrow K \cap L} \times \mathbf{X}_{L \backslash K}$, which means again that $C_{1}=C_{2}$.

Remark: It is not difficult to show that a marginal basic assignment of a cylindrical assignment is again cylindrical. However, it is important to realize that, as we illustrate in the following simple example, an analogous property for sparse basic assignments does not hold. Nevertheless, the main advantage of sparse basic assignments is the fact that the number of their focal elements does not exceed the cardinality of the respective frame of discernment, i.e. the number of probabilities necessary to define a general probability distribution.

Example 3. Consider 2-dimensional case with $\mathbf{X}_{1}=$ $\{a, \bar{a}\}$ and $\mathbf{X}_{2}=\{b, \bar{b}\}$ and basic assignment $m$ in Table 5 . From Table 6 one can immediately see that while marginal basic assignment $m^{\downarrow\{2\}}$ is sparse, the other marginal assignment $m^{\downarrow\{1\}}$ is not.

Remark: Now we are ready to answer the question raised at the beginning of the previous section: what are the basic assignments which are obtained from Bayesian basic assignments by a multiple application of the operator of composition? Since all Bayesian assignments are obviously sparse and cylindrical, Theorems 1 and 2 guarantee that the basic assignments corresponding to compositional models from Bayesian
basic assignments are also cylindrical and sparse. This fact, somehow, justifies the following definition.

Definition 4. Basic assignment is called almost Bayesian if it is sparse and cylindrical.

As said at the beginning of this section, an expressive power of almost Bayesian basic assignments is not too strong. For example, even non-degenerated simple basic assignments are not almost Bayesian. Roughly speaking: Having a Bayesian basic assignment one knows a probability of each point of the frame of discernment. Having an almost Bayesian basic assignment and a fixed point of the frame of discernment one either knows its probability, or knows that it belongs to a cylindrical subset of the frame of discernment among whose elements one cannot make a difference; she knows only the probability of the whole subset. Nevertheless, let us stress once more that the importance of almost Bayesian assignments is in the fact that they describe compositional models constructed from an arbitrary system of low-dimensional probability distributions, which means that even in situations when probabilistic operator of composition is not defined. In this way we are getting a slight extension of probability theory.

## 5 Conditional independence

In this paper our attention is concentrated on properties of basic assignments which are, in a way, promising from the point of view of computational complexity. Last section was devoted to almost Bayesian basic assignment whose number of focal elements is not higher than the number of probabilities by which a general probability distribution must be specified.

It is well known that efficiency of Bayesian models is based on making the best of the dependence structure of the model, i.e. taking advantage of the knowledge of conditional independence relations $[8,9]$ holding for the multidimensional distribution in question. This is because the notion of conditional independence in probability theory is equivalent to the notion of factorization: for probability distribution $P$ variables $X$ and $Z$ are conditionally independent given variable $Y$ iff distribution $P(X, Y, Z)$ is uniquely determined by its marginals $P(X, Y)$ and $P(Y, Z)$. Unfortunately, as shown by Studeny $[8,1]$, the notion of conditional non-interactivity (Shenoy's factorization [7], Studený conditional independence [8]) presented in [1] is not consistent with marginalization: there are situations when for two consistent basic assignments there does not exist their common extension with the respective conditional non-interactivity (for more precise explanation see footnote no. 6).

Therefore, in this paper we are going to eliminate this drawback using the definition of conditional independence for basic assignments introduced in [4], which is in fact based on the notion of factorization. Moreover we will present new proofs showing that for this concept all the semigraphoid axioms hold true. These proofs will be based on the fundamental properties of the operator of composition presented in Lemma 2. It should be stressed that the novelty of these proofs is mainly in application of property (iv) of Lemma 2, which seems to be surprisingly weak (and which, in a way, extends property (ii) of the same lemma).
Let us consider an arbitrary basic assignment. We will say that two groups of variables are conditionally independent given the third group of variables if the respective marginal basic assignment can be decomposed (factorized) in the way that it can be expressed as a composition of its respective smaller marginal assignments. Precisely this notion is introduced in the following definition.

Definition 5. Consider a basic assignment $m$ on $\mathbf{X}_{N}$ and three disjoint index sets $K, L, M \subset N, K \neq \emptyset \neq$ $L$. We say that groups of variables $X_{K}$ and $X_{L}$ are conditionally independent given variables $X_{M}$ if

$$
m^{\downarrow K \cup L \cup M}=m^{\downarrow K \cup M} \triangleright m^{\downarrow L \cup M}
$$

In symbol this fact will be recorded $K \Perp_{m} L \mid M$.

Example 4. Consider a basic assignment $m$ on the same 3-dimensional binary frame of discernment as in previous examples: $\mathbf{X}_{1} \times \mathbf{X}_{2} \times \mathbf{X}_{3}$. If variables $X_{1}$ and $X_{2}$ are independent, i.e. $1 \Perp_{m} 2$, from Definition 1 one can immediately see that for all focal elements $C \subseteq \mathbf{X}_{1} \times \mathbf{X}_{2}$ of the 2-dimensional marginal $m^{\downarrow\{1,2\}}$ it holds that $C=C^{\downarrow\{1\}} \otimes C^{\downarrow\{2\}}$. It means that from all 15 non-empty subsets of $\mathbf{X}_{1} \times \mathbf{X}_{2}$ only 9 of them are potential focal elements (six subsets of $\mathbf{X}_{1} \times \mathbf{X}_{2}$ that cannot be focal elements are listed in Example 1). Naturally, this condition on focal elements is only a necessary condition for the independence. This condition is not sufficient. For example, the reader can easily check that the two basic assignments $m_{1} \triangleright m_{2}$ from Table 2 and $m_{3}$ from Table 3 (both defined on $\mathbf{X}_{1} \times \mathbf{X}_{2}$ ) have the same marginal assignments: $\left(\left(m_{1} \triangleright m_{2}\right)^{\downarrow\{1\}}=m_{3}^{\downarrow\{1\}}=m_{1}\right.$ and $\left.\left(m_{1} \triangleright m_{2}\right)^{\downarrow\{2\}}=m_{3}^{\downarrow\{2\}}=m_{2}\right)$. Moreover, for all of their focal elements the required property $C=C \downarrow\{1\} \otimes C^{\downarrow\{2\}}$ holds true and simultaneously

$$
1 \Perp_{m_{1} \triangleright m_{2}} 2 \quad \text { and } \quad 1 \not \Perp_{m_{3}} 2 .
$$

Analogously to what has just been said about (unconditional) independence, there is a necessary condition
also on focal elements of basic assignments with conditional independence. Conditional independence

$$
1 \Perp_{m} 3 \mid 2
$$

means that all focal elements $C \subseteq \mathbf{X}_{\{1,2,3\}}$ of $m$ must be of the form

$$
C=C^{\downarrow\{1,2\}} \otimes C^{\downarrow\{2,3\}}
$$

It is not difficult to show that this property holds true only for 99 out of all possible 255 nonempty subsets of $\mathbf{X}_{\{1,2,3\}}$.

In the rest of this section we will show that the ternary relation $K \Perp_{m} L \mid M$ is a semigraphoid, i.e. it meets the four semigraphoid axioms listed below. For this, we will exclusively use the properties of the operator of composition presented in Lemma 2. In what follows, each axiom is reformulated into the language of composition and the corresponding theorem is proved.

## Symmetry

$$
I \Perp_{m} J\left|L \Longrightarrow J \Perp_{m} I\right| L
$$

Theorem 3. If $m^{\downarrow I \cup J \cup L}=m^{\downarrow I \cup L} \triangleright m^{\downarrow J \cup L}$ then also $m^{\downarrow I \cup J \cup L}=m^{\downarrow J \cup L} \triangleright m^{\downarrow I \cup L}$.

Proof. The assertion follows immediately from the fact that marginals $m^{\downarrow I \cup L}$ and $m^{\downarrow J \cup L}$ are consistent, and therefore property (iii) may be applied

$$
m^{\downarrow I \cup L} \triangleright m^{\downarrow J \cup L}=m^{\downarrow J \cup L} \triangleright m^{\downarrow I \cup L}
$$

## Decomposition

$$
I \Perp_{m} J \cup K\left|L \Longrightarrow I \Perp_{m} K\right| L
$$

Theorem 4. If $m^{\downarrow I \cup J \cup K \cup L}=m^{\downarrow I \cup L} \triangleright m^{\downarrow J \cup K \cup L}$ then also $m^{\downarrow I \cup K \cup L}=m^{\downarrow I \cup L} \triangleright m^{\downarrow K \cup L}$.

Proof. The assertion will be obtained just by application of properties (iv) and (ii)

$$
\begin{aligned}
m^{\downarrow I \cup K \cup L} & =\left(m^{\downarrow I \cup J \cup K \cup L}\right)^{\downarrow I \cup K \cup L} \\
& =\left(m^{\downarrow I \cup L} \triangleright m^{\downarrow J \cup K \cup L}\right) \downarrow I \cup K \cup L \\
& =\left(\left(m^{\downarrow I \cup L} \triangleright m^{\downarrow K \cup L}\right) \triangleright m^{\downarrow J \cup K \cup L}\right)^{\downarrow I \cup K \cup L} \\
& =m^{\downarrow I \cup L} \triangleright m^{\downarrow K \cup L} .
\end{aligned}
$$

## Weak Union

$I \Perp_{m} J \cup K\left|L \Longrightarrow I \Perp_{m} J\right| K \cup L$
Theorem 5. If $m^{\downarrow I \cup J \cup K \cup L}=m^{\downarrow I \cup L} \triangleright m^{\downarrow J \cup K \cup L}$ then also $m^{\downarrow I \cup J \cup K \cup L}=m^{\downarrow I \cup K \cup L} \triangleright m^{\downarrow J \cup K \cup L}$.

Proof. To prove this assertion we have to realize that, due to property (iv),

$$
m^{\downarrow I \cup L} \triangleright m^{\downarrow J \cup K \cup L}=\left(m^{\downarrow I \cup L} \triangleright m^{\downarrow K \cup L}\right) \triangleright m^{\downarrow J \cup K \cup L},
$$

and that, because the assumptions of Theorem 4 are fulfilled, also

$$
m^{\downarrow I \cup K \cup L}=m^{\downarrow I \cup L} \triangleright m^{\downarrow K \cup L}
$$

Using these two equalities we finish the proof in a simple way

$$
\begin{aligned}
m^{\downarrow I \cup J \cup K \cup L} & =m^{\downarrow I \cup L} \triangleright m^{\downarrow J \cup K \cup L} \\
& =\left(m^{\downarrow I \cup L} \triangleright m^{\downarrow K \cup L}\right) \triangleright m^{\downarrow J \cup K \cup L} \\
& =m^{\downarrow I \cup K \cup L} \triangleright m^{\downarrow J \cup K \cup L} .
\end{aligned}
$$

## Contraction

$I \Perp_{m} K\left|L \& I \Perp_{m} J\right| K \cup L \Longrightarrow I \Perp_{m} J \cup K \mid L$
Theorem 6. If $m^{\downarrow I \cup K \cup L}=m^{\downarrow I \cup L} \triangleright m^{\downarrow K \cup L}$, and $m^{\downarrow I \cup J \cup K \cup L}=m^{\downarrow I \cup K \cup L} \triangleright m^{\downarrow J \cup K \cup L}$, then also $m^{\downarrow I \cup J \cup K \cup L}=m^{\downarrow I \cup L} \triangleright m^{\downarrow J \cup K \cup L}$.

Proof. We will follow the same idea as in the preceding proof but in the reverse order. First, we will use property (iv) and then both assumptions of this assertion.

$$
\begin{aligned}
m^{\downarrow I \cup L} \triangleright m^{\downarrow J \cup K \cup L} & =\left(m^{\downarrow I \cup L} \triangleright m^{\downarrow K \cup L}\right) \triangleright m^{\downarrow J \cup K \cup L} \\
& =m^{\downarrow I \cup K \cup L} \triangleright m^{\downarrow J \cup K \cup L} \\
& =m^{\downarrow I \cup J \cup K \cup L} .
\end{aligned}
$$

## 6 Conclusions

In the paper we dealt with the two problems connected with computational complexity of DempsterShafer theory of evidence. Since full generality of the models leads to exponential grows of space and computational complexity we showed that focusing our attention only to models, which are constructed from Bayesian basic assignments by application of the operator of composition, one does not get beyond the boundaries of a rather limited class of models, which are called in the paper almost Bayesian. The most advantageous characteristics of these models is the fact that though they are able to describe a special type
of an ignorance, they do not have a higher space requirements than classical probabilistic models.
The other goal of this paper was to show that when accepting the notion of conditional independence based on factorization corresponding to the operator of composition, one can easily prove validity of semigraphoid axioms just with the help of the four very elementary properties from Lemma 2. Since the same idea was employed by Prakash P. Shenoy in [7], a very natural question arises: what is the relation of composition introduced in this paper and the Shenoy's notion of combination?
Looking at Shenoy axioms ${ }^{4}$ C1, C2 and C3 we see that Shenoy's axiom C1 (Domain) is equivalent to property (i) of Lemma 2 and therefore it holds also for our composition. However Shenoy's axioms C2 (Associative) and C3 (Commutative) hold for composition only under special conditions. The operator of composition is commutative only for consistent basic assignments; point (iii) of Lemma 2. In definition of conditional independence (Definition 5 of this paper) we consider only composition of consistent assignments (marginals of the considered basic assignment) and therefore we were able to prove axiom of Symmetry. Nevertheless, associativity holds for the operator of composition only under very specific conditions ${ }^{5}$ and therefore the Shenoy's proofs cannot be used. Moreover, property (ii) of Lemma 2 does not hold for Shenoy's combination. So, one cannot be surprised that both of the definitions of conditional independence (i.e. the one proposed in this paper and Shenoy's conditional independence following from the definitions in Section 5 of [7]) are different from each other. They coincide only for unconditional independence and for conditional independence in case of Bayesian basic assignments. Moreover, as we showed in [4], our concept of conditional independence does not suffer from the drawback described in detail in [1], where the authors show that the notion of conditional independence used by Shenoy is not consistent with marginalization ${ }^{6}$. Therefore, we can conclude that our concept of conditional independence seems to meet better some of the intuitive requirements. Nevertheless, a question what is the relation of this notion and concepts of conditional basic assignments remains still open.

[^4]
## Acknowledgements

The research was financially supported by GAČR under the grant no. ICC/08/E010, and 201/09/1891, and by Ministry of Education of the Czech Republic by grants no. 1M0572 and 2C06019.

## References

[1] B. Ben Yaghlane, Ph. Smets, and K. Mellouli, "Belief Function Independence: II. The Conditional Case," Int. J. of Approximate Reasoning, vol. 31, no. (1-2), pp. 31-75, 2002.
[2] R. Jiroušek, "Composition of probability measures on finite spaces," Proc. of the 13th Conf. Uncertainty in Artificial Intelligence UAI'97, (D. Geiger and P. P. Shenoy, eds.). Morgan Kaufmann Publ., San Francisco, California, pp. 274281, 1997.
[3] R. Jiroušek, "On a Conditional Irrelevance Relation for Belief Functions based on the Operator of Composition," Dynamics of Knowledge and Belief (Ch. Beierle, G. Kern-Isberner, eds.) Proceedings of the Workshop at the 30th Annual German Conference on Artificial Intelligence, Fern Universität in Hagen, Osnabrück, 2007, pp.28-41.
[4] R. Jiroušek, J. Vejnarová, "Compositional Models and Conditional Independence in Evidence Theory," submitted to Int. J. of Approximate Reasoning.
[5] R. Jiroušek, J. Vejnarová and M. Daniel, "Compositional models of belief functions," Proc. of the 5th Symposium on Imprecise Probabilities and Their Applications (G. de Cooman, J. Vejnarová and M. Zaffalon, eds.), Charles University Press, Praha, pp. 243-252, 2007.
[6] R. Jiroušek, J. Vejnarová, "There are Combinations and Compositions in Dempster-Shafer Theory of Evidence," submitted to WUPES'09.
[7] P. P. Shenoy, "Conditional independence in valuation-based systems," Int. J. of Approximate Reasoning, vol. 10, no. 3, pp. 203-234, 1994.
[8] M. Studený, "Formal properties of conditional independence in different calculi of AI," Proceedings of European Conference on Symbolic and quantitative Approaches to Reasoning and Uncertainty ECSQARU'93, (K. Clarke, R. Kruse and S. Moral, eds.). Springer-Verlag, 1993, pp. 341351.
[9] M. Studený, "On stochastic conditional independence: the problems of characterization and description," Annals of Mathematics and Artificial Intelligence, vol. 35, p. 323-341, 2002.
[10] J. Vejnarová, "Composition of possibility measures on finite spaces: preliminary results," Proceedings of 7th International Conference on Information Processing and Management of Uncertainty in Knowledge-based Systems IPMU'98, (B. Bouchon-Meunier, R.R. Yager, eds.). Editions E.D.K. Paris, 1998, pp. 25-30.
[11] J. Vejnarová, "Possibilistic independence and operators of composition of possibility measures," Prague Stochastics'98, (M. Hušková , J. Á. Víšek, P. Lachout, eds.) JČMF, 1998, pp. 575-580.


[^0]:    The London
    Mathematical
    Society
    
    http://www.lms.ac.uk/

[^1]:    ${ }^{1}$ In all examples in this paper we record in tables only focal elements. It means that for all subsets of space of discernment which are not included in the respective tables their respective basic assignment equals 0 .

[^2]:    ${ }^{2}$ The simplest example of non-commutativity of the operator of composition can be got by considering two different assignments on the same frame of discernment. Then using property (i) of Lemma 2 we see that their composition is defined on the same frame of discernment as the considered original assignments and the non-commutativity of the operator $\triangleright$ immediately follows from property (ii) of Lemma 2.

[^3]:    ${ }^{3}$ Detailed study of formal similarities of these two operators will appear in [6].

[^4]:    ${ }^{4}$ We do not comment axiom C4 (Zero) because we consider only normalized basic assignments.
    ${ }^{5}$ For example, for basic assignments $m_{1}, m_{2}, m_{3}$ defined on $\mathbf{X}_{K_{1}}, \mathbf{X}_{K_{2}}, \mathbf{X}_{K_{3}}$, respectively
    $K_{1} \supseteq\left(K_{2} \cap K_{3}\right) \quad \Longrightarrow \quad\left(m_{1} \triangleright m_{2}\right) \triangleright m_{3}=m_{1} \triangleright\left(m_{2} \triangleright m_{3}\right)$.
    ${ }^{6}$ Roughly speaking: one can find two consistent basic assignments $m_{1}, m_{2}$, on $\mathbf{X}_{1} \times \mathbf{X}_{2}$ and $\mathbf{X}_{2} \times \mathbf{X}_{3}$, respectively, for which there does not exist a 3-dimensional basic assignment $m$ on $\mathbf{X}_{1} \times \mathbf{X}_{2} \times \mathbf{X}_{3}$ having $m_{1}$ and $m_{2}$ as its marginals, and for which $1 \Perp_{m} 3 \mid 2$.

