On colorings of bivariate random sequences

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Abstract—The ergodic sequences consisting of vectors (ξ_n, η_n) , $n \ge 1$, over a finite alphabet $A \times B$ are colored with $\lfloor e^{n\alpha} \rfloor$ colors for A^n and $\lfloor e^{n\beta} \rfloor$ colors for B^n . Generic behavior of the colorings in terms of probabilities of monochromatic rectangles intersected with typical sets is examined. When n increases a big majority of pairs of colorings produces rectangles whose probabilities are bounded uniformly from above. Limiting rates of bounds are worked out in all regimes of the rates α and β of colorings. As a consequence, generic behavior of the colorings in terms of Shannon entropies of the partitions into rectangles is described.

I. INTRODUCTION

Let N be a finite set and $(\xi_i)_{i \in N}$ a random vector taking a finite number of values. The collection of the Shannon entropies $H(\xi_i, i \in I)$, $I \subseteq N$, of all the subvectors of the vector can be interpreted as an entropic point of a Euclidean space. The last decade has seen renewed investigations of regions of the entropic points and closely related information theoretical inequalities [10], [4], [5], [6], [7]. Motivation has been drawn from numerous schemes of information theory [2], [9] and elsewhere.

Properties of regions of the entropic points and their limits were studied recently in [6] by taking independent identically distributed copies $(\xi_i^{(n)})_{i \in N}$, $n \ge 1$, of the vector and by randomly coloring the copies $\xi_i^{(1)}, \ldots, \xi_i^{(n)}$ of a *single* distinguished variable ξ_i with $\lfloor e^{n\alpha_i} \rfloor$ colors, $\alpha_i \ge 0$. In this way new entropic points were constructed from old ones and their limits were described in [6, Theorem 3].

In this framework, a natural idea is to color independently the copies $\xi_i^{(1)}, \ldots, \xi_i^{(n)}$ of *each* variable ξ_i with $\lfloor e^{n\alpha_i} \rfloor$ colors, $i \in N$, and to investigate the entropies of the partitions into monochromatic blocks. Thus, the *N*-tuples of colorings come into play and a majority of them is expected to have a similar, in some sense generic, behaviour in terms of entropies.

In this contribution, a restriction is made to bivariate vectors. The multivariate case likely differs only by additional technicalities and will be treated elsewhere. Instead of independent copies of a bivariate vector, an ergodic sequence consisting of vectors (ξ_n, η_n) , $n \ge 1$, over a finite alphabet $A \times B$ is considered. The more general assumption of ergodicity entails no additional technical complications.

A k-coloring of a set X is any mapping f of X into the set of colors $\{1, \ldots, k\}$, to be denoted by \hat{k} . The cardinality of X is denoted by |X|. Michal Kupsa

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In the bivariate sequence, first marginal sequence ξ_1, \ldots, ξ_n is colored with $k_n = \lfloor e^{n\alpha} \rfloor$ colors and second one η_1, \ldots, η_n with $\ell_n = \lfloor e^{n\beta} \rfloor$ colors, $\alpha, \beta \ge 0$. If

$$f_n \colon A^n \to \widehat{k_n} \quad \text{and} \quad g_n \colon B^n \to \widehat{\ell_n}$$

are two such colorings, the product $A^n \times B^n$ partitions into the (f_n, g_n) -monochromatic rectangles

$$f_n^{-1}(i) \times g_n^{-1}(j), \qquad i \in \widehat{k_n}, j \in \widehat{\ell_n}.$$

The main result, Theorem 2 in Section IV, deals with uniform upper bounds on probabilities of these rectangles intersected with special typical sets $Z_n \subseteq A^n \times B^n$. The bounds decay exponentially with the rate

$$h_{\xi\eta}^{\alpha\beta} = \min\left\{h_{\xi\eta}, \alpha + h_{\eta}, h_{\xi} + \beta, \alpha + \beta\right\}$$

where $h_{\xi\eta}$, h_{ξ} and h_{η} are the entropy rates of the ergodic sequence and their marginals, respectively. With increasing *n* the bounds become valid for more and more decisive majority of the pairs (f_n, g_n) of colorings, as specified through a notion of a convergence to zero faster than exponentially (f.t.e.).

Using Theorem 2, the generic behaviour of the colorings in terms of entropies of partitions is described in Corollary 2: with increasing n for majority of the pairs (f_n, g_n) of colorings the Shannon entropy of the partition into the (f_n, g_n) -monochromatic rectangles is lower bounded by a sequence na_n such that a_n converges to $h_{\xi\eta}^{\alpha\beta}$. This rate is the highest possible, see Remark 2.

Main tools that underly proofs are collected in Section II. They include a couple of lemmas that go in the spirit back to [1], [3], and their dynamical versions. On the way to the main result, a general asymptotic scheme is singled out in Section III. It comprises crucial arguments towards the proof of Theorem 2 presented in Section IV.

II. PRELIMINARIES

The following lemma is a minor generalization and reformulation of [6, Lemma 6, p. 322]. The starting idea goes back to [1, Lemma 3.1, p. 230], see [6, Remark 4, p. 323]. Results of this sort have been of importance when studying common randomness and secrecy capacities, see [3, Appendix B]. **Lemma 1.** Let $k = e^q$ be a positive integer and Q a finite set of measures on a finite set X such that

$$Q(x) \leq e^{-v} \cdot Q(X), \qquad x \in X, Q \in \mathcal{Q},$$
 (1)

for some v. If $w \leq q$ then the number of those k-colorings f of X that violate

$$Q(f^{-1}(i)) \leqslant e^{-w} \cdot Q(X), \qquad i \in \widehat{k}, Q \in \mathcal{Q},$$
 (2)

is at most

$$|X|^k \cdot k |\mathcal{Q}| \exp\left[-\frac{1}{2}e^{v-q} \left(e^{q-w}-1\right)(q-w)\right].$$

Proof: When Q is empty or contains only the zero measure then (2) is never violated, and the assertion holds trivially. Otherwise, when Q consists of probability measures this is a reformulation of [6, Lemma 6] where $\varepsilon = e^{q-w} - 1$ is nonnegative because $w \leq q$. The general case follows directly from this special one.

In Lemma 1, some colorings f partition X into monochromatic blocks $f^{-1}(i)$ such that their measures $Q(f^{-1}(i))$ are bounded from above, uniformly in i and Q, while the number of the remaining colorings has an explicit upper bound. A symmetric counterpart of the lemma will be needed in the sequel. For readers convenience, its proof is worked out based on ideas of the proof of [6, Lemma 6].

Lemma 2. Under the assumptions of Lemma 1, if $w \ge q$ then the number of those k-colorings f of X that violate

$$Q(f^{-1}(i)) \ge e^{-w} \cdot Q(X), \qquad i \in \widehat{k}, Q \in \mathcal{Q}, \qquad (3)$$

is at most

$$|X|^k \cdot k |\mathcal{Q}| \exp \left[-\frac{1}{2} e^{v-q} \left(1 - e^{q-w} \right) \ln(2 - e^{q-w}) \right].$$

Proof: Let the vector $(\eta_x)_{x \in X}$ consist of independent random variables, each one distributed uniformly on \hat{k} . Thus, a realization of the vector is a k-coloring of X. For $i \in \hat{k}$ and $Q \in Q$ let ζ_x be equal to -(k-1)Q(x) if $\eta_x = i$ and to Q(x)otherwise, $x \in X$. Then, these variables are independent and centered, that is their expectations are equal to zero. By (1), they are bounded in the absolute value by $e^{q-v}Q(X)$ and the sum of variances $(k-1)\sum_{x \in X}Q(x)^2$ is majorized by $e^{q-v}Q(X)^2$. Hence, for $\varepsilon \ge 0$ the inequality

$$\Pr(\sum_{x \in X} \zeta_x > \varepsilon Q(X)) \leq \exp\left[-\frac{\varepsilon}{2e^{q-v}} \ln(1+\varepsilon)\right]$$

follows from [6, Lemma 5] whenever Q(X) > 0, and holds trivially otherwise. This probability multiplied by $|X|^k$ is equal to the number of k-colorings f that satisfy

$$-\sum_{x\in f^{-1}(i)} (k-1)Q(x) + \sum_{x\in X\setminus f^{-1}(i)} Q(x) > \varepsilon \ Q(X) \,.$$

This inequality rewrites to $kQ(f^{-1}(i)) < (1-\varepsilon)Q(X)$. Since w is at least q the choice $\varepsilon = 1 - e^{q-w}$ is possible. Therefore, the number of those k-colorings f that violate the inequality $Q(f^{-1}(i)) \ge e^{-w} \cdot Q(X)$ is at most

$$|X|^k \cdot \exp\left[-\frac{1}{2}e^{v-q} (1-e^{q-w}) \ln(2-e^{q-w})\right].$$

This bound does not depend on Q and i whence the assertion follows.

Dynamical versions of Lemmas 1 and 2 are prepared below for later purposes.

Lemma 3. For $n \ge 1$ let Q_n be a set of measures on a finite set X_n such that

$$Q(x) \leqslant e^{-nr_n} \cdot Q(X_n), \qquad x \in X_n, Q \in \mathcal{Q}_n, \quad (4)$$

for a sequence r_n that converges to a finite limit h. If $\alpha \ge 0$, $k_n = \lfloor e^{n\alpha} \rfloor$ and $s_n = \min\{r_n, \alpha, h\} - 2n^{-1/2}$ then the proportion of those k_n -colorings f_n of X_n that violate

$$Q(f_n^{-1}(i)) \leqslant e^{-ns_n} \cdot Q(X_n), \qquad i \in \widehat{k_n}, Q \in \mathcal{Q}_n, \quad (5)$$

is at most

$$e^{n\alpha} |\mathcal{Q}_n| \exp\left[-\frac{1}{2}(e^{\sqrt{n}}-1)\right].$$

Proof: If $q_n = \ln k_n$ then $n\alpha \ge q_n \ge n\alpha - 1$, and thus $q_n - ns_n \ge \sqrt{n}$. Lemma 1 implies that the proportion of those k_n -colorings f_n that violate (5) is upper bounded by

$$e^{n\alpha} |\mathcal{Q}_n| \exp\left[-\frac{1}{2}e^{nr_n-q_n}\left(e^{q_n-ns_n}-1\right)\sqrt{n}\right].$$

Omitting \sqrt{n} , the bracket is dominated by

$$-\frac{1}{2}e^{n\min\{r_n,\alpha,h\}-n\alpha} (e^{q_n-ns_n}-1) \leqslant -\frac{1}{2}(e^{\sqrt{n}}-1)$$

whence the assertion follows.

Let us say that a sequence of nonnegative numbers p_n grows at most exponentially if the sequence $\frac{1}{n} \ln p_n$ is bounded from above. The sequence p_n goes to zero faster than exponentially (f.t.e.) if $\frac{1}{n} \ln p_n$ tends to $-\infty$.

Lemma 3 is mostly used in a limiting version that only states existence of a sequence s_n .

Corollary 1. If a sequence Q_n satisfies (4) with some $r_n \to h$, $|Q_n|$ grows at most exponentially, $\alpha \ge 0$ and $k_n = \lfloor e^{n\alpha} \rfloor$ then there exists a sequence s_n converging to $\min\{\alpha, h\}$ such that the proportion of those k_n -colorings f_n of X_n that violate (5) goes to zero f.t.e.

A counterpart of this corollary is needed as well.

Lemma 4. Under the same assumptions as in Corollary 1, if $\alpha < h$ then there exists a sequence t_n converging to α such that the proportion of those k_n -colorings f_n of X_n that violate

$$Q_n(f_n^{-1}(i)) \ge e^{-nt_n} \cdot Q_n(X_n), \qquad i \in \widehat{k_n}, Q_n \in \mathcal{Q}_n,$$
(6)

goes to zero f.t.e.

Proof: Let $q_n = \ln k_n$ and $|Q_n| \leq e^{nu}$ for some u. If $w_n = n\alpha + 1$ then $w_n \geq q_n + 1$ and the sequence $t_n = w_n/n$ converges obviously to α . Hence, Lemma 2 implies that the proportion of those k_n -colorings f_n of X_n that violate (6) is upper bounded by

$$e^{n\alpha} e^{nu} \exp\left[-\frac{1}{2} e^{nr_n - n\alpha} \left(1 - \frac{1}{e}\right) \ln\left(2 - \frac{1}{e}\right)\right].$$

This expression vanishes f.t.e.

III. MAIN ASYMPTOTIC SCHEME

In this section a general limiting scheme is presented that abstracts some typical situations encountered in bivariate ergodic sequences. This is believed to provide a better insight and, at the same time, to simplify and shorten proofs.

Theorem 1. For $n \ge 1$ let Q_n be a finite measure on a finite set $X_n \times Y_n$ such that the cardinalities of the sets

$$\{y \in Y_n \colon Q_n(X_n \times \{y\}) \neq 0\}$$
(7)

grow at most exponentially. Let h_X , h_Y and h_{XY} be numbers such that $h_{XY} \leq h_X + h_Y$, the inequalities

$$Q_n(\{x\} \times Y_n) \leqslant e^{-nr_n} \cdot Q_n(X_n \times Y_n), \quad x \in X_n, \quad (8)$$

$$Q_n(X_n \times \{y\}) \leqslant e^{-ns_n} \cdot Q_n(X_n \times Y_n), \quad y \in Y_n, \quad (9)$$

hold with converging sequences $r_n \rightarrow h_X$, $s_n \rightarrow h_Y$, and

$$Q_n(x,y) \leqslant e^{-nt_n} \cdot Q_n(X_n \times \{y\}), \ (x,y) \in X_n \times Y_n, \ (10)$$

be satisfied with a converging sequence $t_n \to h_{XY} - h_Y$. Let further $\alpha, \beta \ge 0$, $k_n = \lfloor e^{n\alpha} \rfloor$, $\ell_n = \lfloor e^{n\beta} \rfloor$ and a pair (f_n, g_n) consist of a k_n -coloring f_n of X_n and an ℓ_n -coloring g_n of Y_n . If $\alpha < h_X$ or $\beta \ge h_Y$ then there exists a sequence

$$w_n \to h_{XY}^{\alpha\beta} = \min\{h_{XY}, \alpha + h_Y, h_X + \beta, \alpha + \beta\}$$

such that the proportion of those pairs (f_n, g_n) that violate

$$Q_n(f_n^{-1}(i) \times g_n^{-1}(j)) \leqslant e^{-nw_n} \cdot Q_n(X_n \times Y_n),$$

$$i \in \widehat{k_n}, j \in \widehat{\ell_n},$$
(11)

goes to zero f.t.e.

Proof: Let Q_n^X be the family consisting of the single measure on X_n given by $x \mapsto Q_n(\{x\} \times Y_n)$. By (8), Corollary 1 applies to the sequence Q_n^X . There exists a sequence u_n^+ converging to $\min\{\alpha, h_X\}$ such that the proportion of those k_n -colorings f_n of X_n that violate

$$Q_n(f_n^{-1}(i) \times Y_n) \leqslant e^{-nu_n^+} \cdot Q_n(X_n \times Y_n), \quad i \in \widehat{k_n}, \quad (12)$$

goes to zero f.t.e.

Let us consider the set of measures $Q_n^{X|Y}$ on X_n given by

$$x \mapsto Q_n(x, y), \quad y \in Y_n$$

that are nonzero. By (7), the cardinality of this set grows at most exponentially. Hence, Corollary 1 based on (10) implies existence of a sequence p_n^+ converging to $\min\{\alpha, h_{XY} - h_Y\}$ such that the proportion of those k_n -colorings f_n that violate

$$Q_n(f_n^{-1}(i) \times \{y\}) \leqslant e^{-np_n^+} \cdot Q_n(X_n \times \{y\}),$$

$$i \in \widehat{k_n}, y \in Y_n,$$
 (13)

goes to zero f.t.e.

If a coloring f_n satisfies (13) then by (9)

$$Q_n(f_n^{-1}(i) \times \{y\}) \leqslant e^{-n[s_n + p_n^+]} \cdot Q_n(X_n \times Y_n),$$

$$i \in \widehat{k_n}, y \in Y_n,$$
 (14)

where $s_n + p_n^+$ converges to $\min\{\alpha + h_Y, h_{XY}\}$.

1. Let us assume first that $\alpha < h_X$. On account of (8), Lemma 4 applies to the sequence Q_n^X . There exists a sequence u_n^- converging to α such that the proportion of those k_n -colorings f_n that violate

$$Q_n(f_n^{-1}(i) \times Y_n) \ge e^{-nu_n^-} \cdot Q_n(X_n \times Y_n), \quad i \in \widehat{k_n}, \quad (15)$$

goes to zero f.t.e.

Let \mathcal{F}_n be the family of those k_n -colorings f_n that violate an inequality in (12), (13), (14) or (15). By three above convergence statements, the proportion $|\mathcal{F}_n||X_n|^{-k_n}$ goes to zero f.t.e. If $f_n \notin \mathcal{F}_n$ then (14) and (15) combine to

$$Q_n(f_n^{-1}(i) \times \{y\}) \leqslant e^{-nv_n} \cdot Q_n(f_n^{-1}(i) \times Y_n),$$

$$i \in \widehat{k_n}, y \in Y_n,$$
(16)

where $v_n = s_n + p_n^+ - u_n^-$ converges to $\min\{h_Y, h_{XY} - \alpha\}$, denoted in the sequel by h_v .

Let $q_n = \min\{v_n, \beta, h_v\} - 2n^{-1/2}$ play the role of s_n from Lemma 3 that is applied to the sets Q_{f_n} of measures on Y_n given by

$$y \mapsto Q_n(f_n^{-1}(i) \times \{y\}), \quad i \in \widehat{k_n},$$

for k_n -colorings $f_n \notin \mathcal{F}_n$ and $n \ge n_0$. By (16), if \mathcal{G}_{f_n} denotes the family of those ℓ_n -colorings g_n of Y_n that violate

$$Q_n(f_n^{-1}(i) \times g_n^{-1}(j)) \leqslant e^{-nq_n} \cdot Q_n(f_n^{-1}(i) \times Y_n),$$

$$i \in \widehat{k_n}, j \in \widehat{\ell_n}, \quad (17)$$

then the proportion $|\mathcal{G}_{f_n}||Y_n|^{-\ell_n}$ is upper bounded by

$$c_n = e^{n\beta} \left| \mathcal{Q}_{f_n} \right| \exp\left[-\frac{1}{2} \left(e^{\sqrt{n}} - 1 \right) \right].$$

Here, obviously $|\mathcal{Q}_{f_n}| \leq e^{n\alpha}$.

Therefore, the cardinality of the set

$$\mathcal{H}_n = \{ (f_n, g_n) \colon f_n \in \mathcal{F}_n \text{ or } (f_n \notin \mathcal{F}_n \text{ and } g_n \in \mathcal{G}_{f_n}) \}$$

is upper bounded by

$$\mathcal{F}_n ||Y_n|^{\ell_n} + \sum_{f_n \notin \mathcal{F}_n} |\mathcal{G}_{f_n}|$$

and the proportion $|\mathcal{H}_n||X_n|^{-k_n}|Y_n|^{-\ell_n}$ is at most the sum of $|\mathcal{F}_n||X_n|^{-k_n}$ with c_n . It follows that this proportion goes to zero f.t.e.

Let $w_n = q_n + u_n^+$. This sequence converges to

$$\min\{\alpha + \beta, \alpha + h_Y, h_{XY}\}$$

which equals $h_{XY}^{\alpha\beta}$ because $\alpha < h_X$. If a pair of colorings (f_n, g_n) does not belong to \mathcal{H}_n , thus $f_n \notin \mathcal{F}_n$ and $g_n \notin \mathcal{G}_{f_n}$, then (12) and (17) hold. Since (11) is their consequence the proportion of those pairs (f_n, g_n) that violate (11) is upper bounded by $|\mathcal{H}_n||X_n|^{-k_n}|Y_n|^{-\ell_n}$, going to zero f.t.e.

2. It remains to consider $\alpha \ge h_X$ and $\beta \ge h_Y$. Let \mathcal{Q}_n^Y be the family consisting of the single measure on Y_n given by $y \mapsto Q_n(X_n \times \{y\})$. Corollary 1 can be applied to the sequence \mathcal{Q}_n^Y and β in the role of α . By (9) and $\beta \ge h_Y$, there exists a sequence v_n^+ converging to h_Y such that the proportion of those ℓ_n -colorings g_n of Y_n that violate

$$Q_n(X_n \times g_n^{-1}(j)) \leqslant e^{-nv_n^+} \cdot Q_n(X_n \times Y_n), \quad j \in \widehat{\ell_n}, \quad (18)$$

goes to zero f.t.e. If a k_n -coloring f_n satisfies (13) then by the summation of y over $g_n^{-1}(j)$,

$$Q_n(f_n^{-1}(i) \times g_n^{-1}(j)) \leqslant e^{-np_n^+} \cdot Q_n(X_n \times g_n^{-1}(j)),$$

$$i \in \widehat{k_n}, j \in \widehat{\ell_n}.$$
 (19)

Here, $p_n^+ \to h_{XY} - h_Y$ because $\alpha \ge h_X \ge h_{XY} - h_Y$.

Let $w_n = v_n^+ + p_n^+$. This sequence tends to $h_{XY} = h_{XY}^{\alpha\beta}$. Combining (18) and (19) the inequalities (11) follow. Since they are violated only if (13) or (18) fails the proportion of those pairs (f_n, g_n) that violate (11) goes to zero f.t.e. \blacksquare *Remark* 1. Let $\alpha < h_{XY} - h_Y$, $\beta < h_Y$ and the assumptions of Theorem 1 hold. Using (7) and (10), Lemma 4 applies to the sequence $Q_n^{X|Y}$ from the previous proof and implies existence of a sequence p_n^- converging to α such that the proportion of those k_n -colorings f_n that violate

$$Q_n(f_n^{-1}(i) \times \{y\}) \geqslant e^{-np_n^-} \cdot Q_n(X_n \times \{y\}),$$

$$i \in \widehat{k_n}, y \in Y_n,$$
(20)

goes to zero f.t.e. Using (9), Lemma 4 applies to the sequence Q_n^Y . There exists a sequence v_n^- converging to β such that the proportion of those ℓ_n -colorings g_n that violate

$$Q_n(X_n \times g_n^{-1}(j)) \geqslant e^{-nv_n^-} \cdot Q_n(X_n \times Y_n), \quad j \in \widehat{\ell}_n, \quad (21)$$

goes to zero f.t.e. If f_n satisfies (20) then for any ℓ_n -coloring g_n of Y_n

$$Q_n(f_n^{-1}(i) \times g_n^{-1}(j)) \ge e^{-np_n^-} \cdot Q_n(X_n \times g_n^{-1}(j)),$$

$$i \in \widehat{k_n}, j \in \widehat{\ell_n},$$
(22)

by summing over $y \in g_n^{-1}(j)$. Combining (21) and (22) it follows that the sequence $w_n = p_n^- + v_n^-$ tends to $\alpha + \beta$ and the proportion of those pairs (f_n, g_n) that violate

$$Q_n(f_n^{-1}(i) \times g_n^{-1}(j)) \ge e^{-nw_n} \cdot Q_n(X_n \times Y_n), i \in \widehat{k_n}, j \in \widehat{\ell_n}$$

goes to zero f.t.e. This symmetric counterpart of Theorem 1 is interesting per se but not used below. The assumption (8) was not needed.

IV. ERGODIC BIVARIATE SEQUENCES

Let (ξ_n, η_n) , $n \ge 1$, be a bivariate ergodic sequence with the states in a finite product $A \times B$. The distribution of the first n vectors of the sequence is denoted by P_n and its marginals to A^n and B^n by P_n^{ξ} and P_n^{η} , respectively. In this section Theorem 1 is applied to the restrictions of P_n to certain subsets of $X_n \times Y_n = A^n \times B^n$ that are constructed by a notion of entropic typicality. The role of h_{XY} is played by the entropy rate $h_{\xi\eta} = \lim \frac{1}{n} \ln H(P_n)$ of the bivariate sequence, and the entropies rates of the marginal sequences, $h_{\xi} = \lim \frac{1}{n} \ln H(P_n^{\xi})$ and $h_{\eta} = \lim \frac{1}{n} \ln H(P_n^{\eta})$, correspond to h_X and h_Y , respectively.

The following theorem asserts rigorously what was earlier mentioned as the generic behavior of pairs of colorings in terms of probabilities of monochromatic rectangles intersected with typical sets. **Theorem 2.** For $n \ge 1$ let $k_n = \lfloor e^{n\alpha} \rfloor$ and $\ell_n = \lfloor e^{n\beta} \rfloor$ where $\alpha, \beta \ge 0$. There exist a sequence of sets $Z_n \subseteq A^n \times B^n$ and a sequence of numbers w_n such that $P_n(Z_n) \to 1$, $w_n \to h_{\xi\eta}^{\alpha\beta}$ and the proportion of those pairs (f_n, g_n) , consisting of a k_n -coloring f_n of A^n and an ℓ_n -coloring g_n of B^n , that violate

$$P_n\left(\left(f_n^{-1}(i) \times g_n^{-1}(j)\right) \cap Z_n\right) \leqslant e^{-nw_n}, i \in k_n, \ j \in \ell_n, \ (23)$$

goes to zero f.t.e.

Proof: By exchanging the coordinate variables ξ_n and η_n if necessary there is no loss of generality when assuming $\alpha < h_{\xi}$ or $\beta \ge h_{\eta}$. The subset Z_n is constructed below via the entropy-typical sets

$$T^{n}_{\xi,\delta} = \left\{ x \in A^{n} \colon P^{\xi}_{n}(x) \asymp e^{-n[h_{\xi} \pm \delta]} \right\},$$

$$T^{n}_{\eta,\delta} = \left\{ y \in B^{n} \colon P^{\eta}_{n}(y) \asymp e^{-n[h_{\eta} \pm \delta]} \right\},$$

$$T^{n}_{\xi\eta,\delta} = \left\{ (x,y) \in A^{n} \times B^{n} \colon P_{n}(x,y) \asymp e^{-n[h_{\xi\eta} \pm \delta]} \right\}.$$

Here, the symbol \asymp means that the number on the left is between the two numbers on the right.

Since the sequence (ξ_n, η_n) and two marginal sequences are ergodic the Shannon-McMillan-Breiman theorem, known also as the asymptotic equipartition property [8, Sections 1.5 and 1.6], implies that for every positive δ the probabilities $P_n^{\xi}(T_{\xi,\delta}^n)$, $P_n^{\eta}(T_{\eta,\delta}^n)$ and $P_n(T_{\xi\eta,\delta}^n)$ tend to 1. By standard diagonal arguments, there exists a positive sequence $(\delta_n)_{n\geq 1}$ converging to zero, perhaps rather slowly, such that each of the sequences $P_n^{\xi}(T_{\xi,\delta_n}^n)$, $P_n^{\eta}(T_{\eta,\delta_n}^n)$ and $P_n(T_{\xi\eta,\delta_n}^n)$ converges to one.

Let U_n denote the intersection of $T^n_{\xi,\delta_n} \times B^n$, $A^n \times T^n_{\eta,\delta_n}$ and $T^n_{\xi\eta,\delta_n}$. If $p_n = 1 - [1 - P_n(U_n)]^{1/2}$ and

$$Y_n^* = \left\{ y \in B^n \colon P_n((A^n \times \{y\}) \cap U_n) \ge p_n \cdot P_n^\eta(y) \right\}$$

then the estimations

$$p_n[1 - P_n^{\eta}(Y_n^*)] = \sum_{y \in B^n \setminus Y_n^*} p_n \cdot P_n^{\eta}(y)$$

$$\geq \sum_{y \in B^n \setminus Y_n^*} P_n((A^n \times \{y\}) \cap U_n)$$

$$= P_n(U_n) - P_n((A^n \times Y_n^*) \cap U_n)$$

$$\geq P_n(U_n) - P_n^{\eta}(Y_n^*)$$

and $p_n < 1$ imply

$$P_n^{\eta}(Y_n^*) \ge \frac{P_n(U_n) - p_n}{1 - p_n} = p_n \,.$$

Obviously, $P_n^{\eta}(Y_n^*) = 1$ if $p_n = 1$.

Let Z_n denote the intersection of $A^n \times Y_n^*$ with U_n . Then it is possible to conclude subsequently that the sequences $P_n(U_n)$, p_n , $P_n^{\eta}(Y_n^*)$ and $P_n(Z_n)$ converge to one. Let Q_n denote the restriction of P_n to Z_n . To apply Theorem 1, its assumptions are verified as follows.

The set in (7) is contained in $Y_n = B^n$ so that its cardinality grows at most exponential.

Let r_n be equal to $h_{\xi} - \delta_n + \frac{1}{n} \ln P_n(Z_n)$ provided $P_n(Z_n)$ is positive. If $x \notin T^n_{\xi,\delta_n}$ then $Q_n(\{x\} \times Y_n)$ vanishes and the inequality in (8) holds trivially. Otherwise, if $x \in T^n_{\xi,\delta_n}$ then

$$Q_n(\{x\} \times B^n) \leqslant P_n(\{x\} \times B^n) = P_n^{\xi}(x) \leqslant e^{-n[h_{\xi} - \delta_n]}$$
$$= e^{-nr_n} \cdot P_n(Z_n) = e^{-nr_n} \cdot Q_n(A^n \times B^n)$$

so that (8) is verified with a sequence r_n converging to h_{ξ} .

For s_n defined through $h_\eta - \delta_n + \frac{1}{n} \ln P_n(Z_n)$, a verification of (9) is analogous to that of (8) and omitted here.

Let t_n be equal to $h_{\xi\eta} - h_\eta - 2\delta_n + \frac{1}{n}\ln p_n$ provided p_n is positive. If $(x, y) \notin Z_n$ then $Q_n(x, y)$ vanishes and the inequality in (10) holds trivially. Otherwise, $(x, y) \in T^n_{\xi\eta,\delta_n}$ and $y \in T^n_{\eta,\delta_n}$ imply

$$Q_n(x,y) = P_n(x,y) \leqslant e^{-n[h_{\xi\eta} - \delta_n]}$$
$$\leqslant e^{-n[h_{\xi\eta} - \delta_n]} e^{n[h_{\eta} + \delta_n]} P_n^{\eta}(y) \,.$$

Using $y \in Y_n^*$ and $Z_n = (A^n \times Y_n^*) \cap U_n$,

$$p_n \cdot P_n^{\eta}(y) \leqslant P_n((A^n \times \{y\}) \cap U_n) = Q_n(A^n \times \{y\}).$$

Combining above estimations, it follows that (10) is verified with a sequence $t_n \rightarrow h_{\xi\eta} - h_{\eta}$. Obviously, $h_{\xi\eta} \leq h_{\xi} + h_{\eta}$.

Therefore, Theorem 1 implies existence of a sequence v_n converging to $h_{\xi\eta}^{\alpha\beta}$ such that the proportion of those pairs (f_n, g_n) that violate

$$Q_n(f_n^{-1}(i) \times g_n^{-1}(j)) \leqslant e^{-nv_n} \cdot P_n(Z_n), \quad i \in \widehat{k_n}, j \in \widehat{\ell_n},$$

goes to zero f.t.e. Writing $w_n = v_n - \frac{1}{n} \ln P_n(Z_n)$, the above inequalities coincide with (23) and $w_n \to h_{\xi\eta}^{\alpha\beta}$.

The following consequence of Theorem 2 describes what was alluded to as the generic behavior of the colorings in terms of Shannon entropies of the partitions into rectangles.

Corollary 2. If $H(P_n|f_n, g_n)$ denotes the Shannon entropy under P_n of the partition of $A^n \times B^n$ into the (f_n, g_n) -monochromatic rectangles $f_n^{-1}(i) \times g_n^{-1}(j)$, $i \in \widehat{k_n}$, $j \in \widehat{\ell_n}$, then there exists a sequence a_n converging to $h_{\xi\eta}^{\alpha\beta}$ such that the proportion of those pairs (f_n, g_n) of colorings that violate $\frac{1}{n}H(P_n|f_n, g_n) \ge a_n$ goes to zero f.t.e.

Proof: If $Z_n^c = (A^n \times B^n) \setminus Z_n$ then

$$\sum_{D} \left[P_n(D \cap Z_n) \ln \frac{P_n(D \cap Z_n)}{P_n(D)P_n(Z_n)} + P_n(D \cap Z_n^c) \ln \frac{P_n(D \cap Z_n^c)}{P_n(D)P_n(Z_n^c)} \right]$$

is nonnegative by convexity. Here, the summation runs over the (f_n, g_n) -monochromatic rectangles D. This implies

$$H(P_n|f_n,g_n) + \ln 2 \ge -\sum_D P_n(D \cap Z_n) \ln P_n(D \cap Z_n).$$

By Theorem 2, for a sequence w_n converging to $h_{\xi\eta}^{\alpha\beta}$, the summand on the right is majorized by $-nw_nP_n(D\cap Z_n)$. Taking $a_n = w_nP_n(Z_n) - \frac{1}{n}\ln 2$ the assertion follows. *Remark* 2. If the assertion of Corollary 2 holds for a sequence a_n converging to some a instead of $h_{XY}^{\alpha\beta}$ then a cannot exceed this number. This follows from

$$\limsup_{n \to \infty} \sup_{f_n, g_n} \frac{1}{n} H(P_n | f_n, g_n) \leqslant h_{X, Y}^{\alpha, \beta}.$$
(24)

To prove the inequality, $H(P_n|f_n, g_n)$ is majorized by the sum of $H(P_n^{\xi}|f_n)$ and $H(P_n^{\eta}|g_n)$, defined analogously. The former summand is dominated by $H(P_n^{\xi})$ and $n\alpha$ because the partition of A^n into $f_n^{-1}(i)$, $i \in \widehat{k_n}$, has at most k_n blocks. Similarly, the latter summand is dominated by $H(P_n^{\eta})$ and $n\beta$. It follows that the left-hand side of (24) is at most

to now s that the left-hand side of (2+) is at the

$$\min\{h_{\xi},\alpha\} + \min\{h_{\eta},\beta\}.$$

This and $H(P_n|f_n, g_n) \leq H(P_n)$ imply (24).

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