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A Counterexample on Sample-Path Optimality in Stable Markov Decision Chains with the Average Reward Criterion

Rolando Cavazos-Cadena \cdot Raúl Montes-de-Oca \cdot Karel Sladký

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Abstract This note deals with Markov decision chains evolving on a denumerable state space. Under standard continuity-compactness requirements, an explicit example is provided to show that, with respect to a strong sample-path average reward criterion, the Lyapunov function condition does not ensure the existence of an optimal stationary policy.

Keywords Strong sample-path optimality · Lyapunov function condition · Stationary policy · Expected average reward criterion

1 Introduction

This work is concerned with discrete-time Markov decision processes (MDPs) with denumerable state space, compact action sets, and endowed with a long-run average criterion. It is assumed that the reward function and the transition law depend continuously on the applied action, and that the so-called Lyapunov function condition holds. Within this framework, the optimal expected average reward does not depend

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J Optim Theory Appl (2014) 163:674-684

on the initial state and is characterized in terms of the corresponding optimality equation, whose solution renders an *expected* average optimal stationary policy; see [1]. The expected average reward index arises naturally when the controller runs the un-

derlying dynamical system a large number of times, but such an expected criterion does not look appropriate if the system is going to be observed just a few times. In this latter case, it is natural to analyze the behavior of a control policy from a *sample-path perspective*, and the following notion will be used in this note:

A policy π^* is average optimal in the sample-path sense if there exists a constant, say g^* , such that under the action of π^* and regardless the initial state, the average of the observed rewards over a finite horizon *t* converges to g^* as $t \to \infty$ with probability 1, whereas under any other policy the superior limit of such averages is always bounded above by g^* almost surely.

Under the mild continuity-compactness conditions in this note, it was recently shown in [2] that if the existence of a Lyapunov function is complemented with an additional requirement, which is described in Sect. 3, then an expected average optimal stationary policy obtained from the optimality equation is also sample-path average optimal in the sense described above, a result that naturally leads to consider the following question:

• Can the existence of a sample-path optimal stationary policy be ensured under the sole assumption that the decision model admits a Lyapunov function?

The main objective of this work is to exhibit an explicit example showing that the answer to this question is negative, a conclusion that establishes an interesting contrast between the expected and sample-path average criteria. *The organization* of the subsequent material is as follows. In Sect. 2 the decision model and the Lyapunov function condition are briefly discussed, whereas in Sect. 3 the notion of (strong) sample-path average optimal policy and the main question considered in this note are formally stated. Next, in Sect. 4 an MDP admitting a Lyapunov function and satisfying the standard continuity-compactness requirements is introduced and, for such a particular model, it is shown that there is not any stationary policy which is sample-path average optimal in the strong sense of Definition 3.1. Finally, the exposition concludes in Sect. 5 with some brief comments.

2 Decision Model

Throughout the remainder \mathbb{N} stands for the set of all nonnegative integers and the indicator function of an event *A* is denoted by *I*[*A*]; given sequence of events {*A_n*}, the corresponding superior limit is denoted by [*A_n* i.o.], that is, [*A_n* i.o.] := $\bigcap_{m=1}^{\infty} \bigcup_{k=m}^{\infty} A_k$, whereas the class of all real-valued and continuous functions defined on a topological space \mathbb{K} is denoted by $\mathcal{C}(\mathbb{K})$. Now, let $\mathcal{M} = (S, A, \{A(x)\}_{x \in S}, R, P)$ be an MDP, where the state space *S* is a denumerable set endowed with the discrete topology, the action set *A* is a metric space and, for each $x \in S$, $A(x) \subset A$ is the nonempty subset of admissible actions at *x*, whereas $R \in \mathcal{C}(\mathbb{K})$ is the reward function, where $\mathbb{K} := \{(x, a) | x \in S, a \in A(x)\}$ is the space of admissible pairs. On the other hand, $P = [p_{xy}(\cdot)]$ is the controlled transition law on *S* given \mathbb{K} , that is, for all $(x, a) \in \mathbb{K}$ and $y \in S$, the relations $p_{xy}(a) \ge 0$ and $\sum_{y \in S} p_{xy}(a) = 1$ are satisfied. This model \mathcal{M} is interpreted as follows: At each time $t \in \mathbb{N}$ the decision maker knows the previous states and actions and observes the current state, say $X_t = x \in S$. Using that information, the controller selects an action (control) $A_t = a \in A(x)$ and two things happen: a reward R(x, a) is obtained by the controller, and the system moves to a new state $X_{t+1} = y \in S$ with probability $p_{xy}(a)$.

Assumption 2.1

- (i) For each $x \in S$, A(x) is a compact subset of A.
- (ii) For every $x, y \in S$, the mappings $a \mapsto R(x, a)$ and $a \mapsto p_{xy}(a)$ are continuous in $a \in A(x)$.

Policies A policy π is a (measurable) rule for choosing actions which, at each time $t \in \mathbb{N}$, may depend on the current state and on the record of previous states and actions; see, for instance, [3] for details. The class of all policies is denoted by \mathcal{P} and, given the initial state $x \in S$ and the policy π being used for choosing actions, the distribution of the state-action process $\{(X_t, A_t)\}$ is uniquely determined; such a distribution is denoted by P_x^{π} , whereas E_x^{π} stands for the corresponding expectation operator. Next, define $\mathbb{F} := \prod_{x \in S} A(x)$ and notice that \mathbb{F} is a compact metric space, which consists of all functions $f: S \to A$ such that $f(x) \in A(x)$ for each $x \in S$. A policy π is *stationary* iff there exists $f \in \mathbb{F}$ such that the equality $A_t = f(X_t)$ is always valid under π ; in this case π and f are naturally identified and, with this convention, $\mathbb{F} \subset \mathcal{P}$.

Expected Average Criterion and Lyapunov Function Condition Assume that $R(X_t, A_t)$ has finite expectation with respect to every distribution P_x^{π} . The (long-run superior limit) expected average reward criterion corresponding to $\pi \in \mathcal{P}$ at state $x \in S$ is defined by

$$J(x,\pi) := \limsup_{k \to \infty} \frac{1}{k} E_x^{\pi} \left[\sum_{t=0}^{k-1} R(X_t, A_t) \right],\tag{1}$$

whereas the corresponding optimal value function is

$$J^*(x) := \sup_{\pi \in \mathcal{P}} J(x, \pi), \quad x \in S;$$
⁽²⁾

a policy $\pi^* \in \mathcal{P}$ is (expected) average optimal if $J(x, \pi^*) = J^*(x)$ for every $x \in S$. A fundamental instrument to analyze the above criterion is the following *optimality equation*:

$$g + h(x) = \sup_{a \in A(x)} \left[R(x, a) + \sum_{y \in S} p_{xy}(a)h(y) \right], \quad x \in S,$$
(3)

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where $g \in \mathbb{R}$ and $h \in \mathcal{C}(S)$ is a given function. Suppose that the pair $(g, h(\cdot))$ satisfies (3) and that the following properties are valid: For each $x \in S$ and $\pi \in \mathcal{P}$,

(i)
$$E_x^{\pi}[|h(X_n)|] < \infty$$
 for each $n \in \mathbb{N}$, and $E_x^{\pi}[|h(X_n)|]/n \to 0$ as $n \to \infty$;

(ii) The mapping
$$a \mapsto \sum_{y \in S} p_{xy}(a)h(y), \ a \in A(x)$$
 is continuous. (4)

Using Assumption 2.1, these requirements yield that

- (a) $J^*(x) = g$ for each $x \in S$, and
- (b) There exists a stationary policy $f \in \mathbb{F}$ satisfying

$$g + h(x) = R(x, f(x)) + \sum_{y \in S} p_{xy}(f(x))h(y), \quad x \in S,$$
 (5)

and such a stationary policy f is average optimal [1, 2].

The existence of a solution $(g, h(\cdot))$ of the optimality equation satisfying the properties (i) and (ii) in (4) and, consequently, rendering the above conclusions (a) and (b), requires some connectedness condition [4]. The following is a general requirement in this direction.

Assumption 2.2 (Lyapunov Function Condition [1]) There exists $z \in S$ and a function $\ell: S \to [1, \infty)$ satisfying the properties (i)–(iii) below:

- (i) $1 + |R(x, a)| + \sum_{y \neq z} p_{xy}(a)\ell(y) \le \ell(x)$ for all $(x, a) \in \mathbb{K}$; (ii) For each $x \in S$, $a \mapsto \sum_{y} p_{xy}(a)\ell(y)$ is a continuous function of $a \in A(x)$;
- (iii) For every $f \in \mathbb{F}$ and $x \in S$, $E_x^f [\ell(X_n)I[T > n]] \to 0$ as $n \to \infty$, where

$$T := \min\{n > 0 | X_n = z\}$$

is the first return time to state z.

Given a state $z \in S$, a function ℓ satisfying the conditions (i)–(iii) in the above assumption is referred to as a *Lyapunov function* for the model \mathcal{M} .

Lemma 2.1 [1] Under Assumptions 2.1 and 2.2, assertions (i)–(iii) below hold:

- (i) $E_x^{\pi}[|R(X_n, A_n)|]$ is finite for every $x \in S$ and $\pi \in \mathcal{P}$;
- (ii) There exists a pair $(g, h(\cdot)) \in \mathbb{R} \times C(S)$ satisfying the optimality equation (3) as well as the two conditions in (4);
- (iii) There exists a policy $f \in \mathbb{F}$ satisfying (5) and such a stationary policy is expected average optimal; moreover, $J^*(\cdot) = g$.

3 Sample-Path Optimality

The expected average criterion in (1) is quite appropriate if the controller repeats the underlying random dynamical experiment many times under similar conditions, but not for a single trial. In this latter case, it is interesting to study the average reward from a *sample-path* point of view.

Definition 3.1 A policy $\pi^* \in \mathcal{P}$ is *(strong)* sample-path average optimal with optimal value $g^* \in \mathbb{R}$ if the following conditions (i) and (ii) hold:

- (i) For each state $x \in S$, $\lim_{n\to\infty} \frac{1}{n} \sum_{t=0}^{n-1} R(X_t, A_t) = g^* P_x^{\pi^*}$ -a.s., and (ii) For every $\pi \in \mathcal{P}$ and $x \in S$, $\limsup_{n\to\infty} \frac{1}{n} \sum_{t=0}^{n-1} R(X_t, A_t) \le g^* P_x^{\pi}$ -a.s..

The above notion is stronger than the idea of sample path optimality that was employed in [5], where the second property in the above definition was replaced by the weaker requirement $\liminf_{n\to\infty}\sum_{t=0}^{n-1} R(X_t, A_t)/n \le g^* P_x^{\pi}$ -a.s.. With respect to this weaker notion of sample-path optimality, it was proved in Theorem 4.1 of the aforementioned paper that the policy $f \in \mathbb{F}$ satisfying (5) is sample-path optimal with optimal sample-path average reward $g^* = J^*(\cdot)$. On the other hand, it was recently shown in [2] that, under Assumptions 2.1 and 2.2, if the Lyapunov function ℓ is such that, regardless of the initial state and the policy employed, the expected average reward corresponding to ℓ^{β} is finite for some $\beta > 2$, then the stationary policy f in (5) is sample-path average optimal in the strong sense. In short, the following results are presently available for MDPs satisfying Assumptions 2.1 and 2.2: (a) There exists a stationary policy which is sample-path optimal with respect to a criterion that is *weaker* than the one specified above; and (b) With respect to the strong idea of sample-path optimality in Definition 3.1, the existence of an optimal stationary policy has been established when the Lyapunov function satisfies an additional requirement. These facts naturally lead to the following question:

Are Assumptions 2.1 and 2.2 sufficient to ensure the existence of a stationary policy that is sample-path optimal in the sense of Definition 3.1?

The main contribution of this note consists in showing that the answer to this question is negative, a conclusion that will be established using an explicit example; for additional results on sample-path optimality in other contexts see, for instance, [6] and [7].

4 The Counterexample

In this section an example will be given to show that Assumptions 2.1 and 2.2 do not generally ensure the existence of a sample-path optimal stationary policy in the sense of Definition 3.1.

Example 4.1 Let the state space S and the action set A be the topological subspaces of the real line given by

 $S = \mathbb{N}, \quad A = \{0\} \cup \{1/k | k = 1, 2, 3, \ldots\},\$

and define the sets of admissible actions by

A(0) = A, and $A(x) = \{1\}, x = 1, 2, 3, \dots$

Next, let the transition law be determined by

$$p_{x0}(1) = 1, \quad x = 1, 2, 3, \dots,$$
 (6)

$$p_{00}(0) = 1$$
, and $p_{0x}(1/x) = \frac{1}{x \log(12 + x)} = 1 - p_{00}(1/x)$, $x = 1, 2, ... (7)$

and, finally, let the reward function $R \colon \mathbb{K} \to \mathbb{R}$ be given by

$$R(x, a) = ax, \quad x \in S, \quad a \in A(x), \tag{8}$$

that is, R(0, a) = 0 for every $a \in A(0) = A$, and R(x, 1) = x for every $x \in S$.

It is not difficult to see that the continuity-compactness conditions in Assumption 2.1 are satisfied for the model in the above example, and it will be shown below that Assumption 2.2 also holds. First, using that A(x) is the singleton {1} for each state $x \neq 0$, observe that

$$\mathbb{F} = \{ f_0, f_1, f_2, \ldots \}, \tag{9}$$

where the policies f_k are determined by $f_k(x) = 1$ for every x = 1, 2, 3, ... and $k \in \mathbb{N}$, and

$$f_0(0) = 0, \quad f_k(0) = 1/k, \quad k = 1, 2, 3, \dots$$
 (10)

Proposition 4.1 *Consider the model* \mathcal{M} *in Example* 4.1, *set* z = 0 *and define the function* $\ell: S \to \mathbb{R}$ *as follows:*

$$\ell(0) = 1 + \frac{2}{\log(12)}, \quad and \quad \ell(x) = x + 1, \quad x = 1, 2, 3, \dots$$
 (11)

With this notation, the following assertions (a) and (b) hold:

- (a) The mapping ℓ is a Lyapunov function for the model \mathcal{M} ;
- (b) The optimal expected average reward is

$$g = \frac{1}{\log(16) + 1/4},\tag{12}$$

and f_4 is the unique (expected) average optimal stationary policy; see (10).

Proof (a) The three requirements in Assumption 2.2 will be verified.

(i) Note that (6)–(8) and (11) together yield that

$$1 + |R(x, 1)| + \sum_{y \neq 0} p_{xy}(1)\ell(y) = 1 + x = \ell(x), \quad x = 1, 2, 3, \dots,$$

$$1 + |R(0, 0)| + \sum_{y \neq 0} p_{0y}(0)\ell(y) = 1 \le \ell(0), \text{ and}$$

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$$1 + |R(0, 1/k)| + \sum_{y \neq 0} p_{0y}(1/k)\ell(y) = 1 + \frac{k+1}{k\log(12+k)}$$
$$\leq 1 + \frac{2}{\log(12)} = \ell(0), \quad k = 1, 2, 3, \dots;$$

from the specification of the action sets, these relations yield that the first condition in Assumption 2.2 is satisfied by ℓ .

(ii) Since A(x) is a singleton for $x \neq 0$ and zero is the unique accumulation point of $A(0) = \{0\} \cup \{1/k | k = 1, 2, 3, ...\}$, it is sufficient to show that

$$\lim_{k \to \infty} \sum_{y \in S} p_{0y}(1/k)\ell(y) = \sum_{y \in S} p_{0y}(0)\ell(y) = \ell(0),$$

a convergence that is valid, since (6), (7) and (11) together imply that

$$\sum_{y \in S} p_{0y}(1/k)\ell(y) = \frac{k+1}{k\log(12+k)} + \left(1 - \frac{1}{k\log(12+k)}\right)\ell(0), \quad k = 1, 2, 3, \dots$$

(iii) The specification of the transition law yields that, for every $x \in S$ and $f \in \mathbb{F}$, the inequality $T \leq 2$ holds P_x^f -a.s., so that $E_x^f[\ell(X_n)I[T > n]] = 0$ for $n \geq 2$.

(b) Using the existence of a Lyapunov function established in the previous part, Lemma 2.1 yields that the optimal expected average reward is constant, say g, and that

$$g = \sup_{f \in \mathbb{F}} J(0, f) = \sup \{ J(0, f_k) | k = 0, 1, 2, 3, \dots \},$$
(13)

where (9) was used to set the second equality. Now, observe that under f_0 the state 0 is absorbing and then, since R(0, 0) = 0, it follows that

$$J(0, f_0) = 0. (14)$$

Suppose now that the system is driven by policy f_k with k > 0. In this case, the specification of the transition law yields that $X_1 = 0$ or $X_1 = k$ with probability 1, and the following assertions hold:

- (i) On the event $[X_1 = 0]$ the equalities T = 1 and $\sum_{t=0}^{T-1} R(X_t, A_t) = 0$ are valid $P_0^{f_k}$ -a.s.;
- (ii) On $[X_1 = k]$, the events [T = 2] and $[\sum_{k=0}^{T-1} R(X_k, A_k) = 0 + k = k]$ occur $P_0^{f_k}$ -a.s.

These facts and the relations $P_0^{f_k}[X_1 = k] = 1/[k \log(12 + k)] = 1 - P_0^{f_k}[X_1 = 0]$ together yield, via the theory of renewal-reward processes [8], that

$$J(0, f_k) = \frac{E_0^{f_k} [\sum_{k=0}^{T-1} R(X_k, A_k)]}{E_0^f [T]}$$
$$= \frac{k(1/[k \log(12+k)])}{(1-1/[k \log(12+k)]) + 2(1/[k \log(12+k)])}$$

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J Optim Theory Appl (2014) 163:674-684

$$= \frac{1/\log(12+k)}{1+1/[k\log(12+k)]}$$
$$= \frac{1}{\log(12+k)+1/k} \le \frac{1}{\log(16)+1/4} = J(0, f_4),$$

where the inequality is strict for $k \neq 4$. The conclusion follows combining this last display with (13) and (14).

If $(g, h(\cdot))$ is a solution of the optimality equation for the MDP in Example 4.1, then f_4 is the unique stationary policy satisfying (5), where the optimal expected average reward g is given in (12). Also, notice that the standard ergodic theorem for Markov chains yields that, for every initial state $x \in S$ and k = 0, 1, 2, 3, ... (see [8]),

$$\lim_{n \to \infty} \frac{1}{n+1} \sum_{k=0}^{n} R(X_t, A_t) = J(0; f_k) \le g = \frac{1}{\log(16) + 1/4} < \frac{1}{2}, \quad P_x^{f_k} \text{-a.s.}$$
(15)

This relation will be used to show that a sample-path average optimal stationary policy in the sense of Definition 3.1 does not exist. To achieve this goal, it is convenient to introduce some notation: For each $t \in \mathbb{N}$ let N_t be the number of visits to state 0 up to time t, i.e.,

$$N_t(X_0, A_0, \dots, X_{t-1}, A_{t-1}, X_t) \equiv N_t \colon = \sum_{k=0}^t I[X_k = 0].$$
(16)

Using that $A(w) = \{1\}$ when $w \in S \setminus \{0\}$, the Markov property yields that, for every $x \in S$, $\delta \in \mathcal{P}$ and $t \in \mathbb{N}$,

$$P_x^{\delta} \big[X_t \neq 0, X_{t+1} \neq 0 | X_0, A_0, \dots, X_t \big] = I[X_t \neq 0] \sum_{y \in S \setminus \{0\}} p_{X_t y}(1) = 0,$$

where (6) was used to set the second equality. Thus,

$$P_x^{\delta}[X_t \neq 0, X_{t+1} \neq 0] = 0, \tag{17}$$

i.e., $P_x^{\delta}[X_t = 0 \text{ or } X_{t+1} = 0] = 1$, an equality that via (16) yields the following conclusion.

Proposition 4.2 *For each initial state* $x \in S$ *and* $\delta \in \mathcal{P}$ *,*

$$P_x^{\delta}[N_t \ge [t/2]] = 1, \quad t = 0, 1, 2, 3, \dots$$

Next, let the policy $\pi \in \mathcal{P}$ be determined as follows: At each time $t \in \mathbb{N}$, under the action of π the control A_t is given by

$$A_t = 1/N_t$$
 if $X_t = 0$, and $A_t = 1$ if $X_t \neq 0$. (18)

Proposition 4.3 *The policy* π *determined by* (18) *satisfies the following properties: For each* $x \in S$,

(i) $P_x^{\pi}[X_t \neq 0 \ i.o.] = 1$,

and then

(ii) $P_x^{\pi}[X_t = 0, X_{t+1} \neq 0 \text{ i.o.}] = 1.$

Proof For each integer $m \in \mathbb{N}$, the specification of the policy π yields that

$$P_x^{\pi} \left[\bigcap_{k=m}^{\infty} [X_k = 0] | X_s, 0 \le s \le m \right]$$

= $I[X_m = 0] \prod_{k=0}^{\infty} p_{00} (1/[k+N_m])$
= $I[X_m = 0] \prod_{k=0}^{\infty} \left(1 - \frac{1}{(k+N_m)\log(12+k+N_m)} \right) = 0.$

where the last equality is due to the relation $\sum_{r=1}^{\infty} 1/[r \log(12+r)] = \infty$. Thus,

$$P_x^{\pi} \left[\bigcap_{k=m}^{\infty} [X_k = 0] \right] = 0, \quad x \in S, \quad m = 0, 1, 2, 3, \dots,$$

a property that is equivalent to the first conclusion. Next, note that

$$P_x^{\pi} \left[[X_{t+1} \neq 0] \setminus [X_t = 0, X_{t+1} \neq 0] \right] = P_x^{\pi} [X_t \neq 0, X_{t+1} \neq 0] = 0,$$

where the second equality is due to (17); so, assertion (ii) follows from part (i).

To conclude, the two previous propositions will be used to show that the MDP in Example 4.1 does not admit a sample path optimal stationary policy in the sense of Definition 3.1.

Proposition 4.4 In Example 4.1 the assertions (i) and (ii) below are valid:

(i) The policy π in (18) satisfies the following property: For each $x \in S$,

$$\limsup_{t\to\infty}\frac{\sum_{k=0}^{t-1}R(X_k,A_k)}{t}\geq \frac{1}{2}\quad P_x^{\pi}\text{-a.s.}$$

Consequently,

(ii) A sample-path average optimal stationary policy does not exist.

Proof (i) Let $x \in S$ be arbitrary and observe that (7) and (18) together imply that

$$P_x^{\pi} [X_t = 0, X_{t+1} \neq 0 | X_0, X_1, \dots, X_t] = I[X_t = 0] \sum_{y \neq 0} p_{0y}(1/N_t)$$

= $I[X_t = 0] p_{0N_t}(1/N_t)$
= $P_x^{\pi} [X_t = 0, X_{t+1} = N_t | X_0, X_1, \dots, X_t],$

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so that $P_x^{\pi}[X_t = 0, X_{t+1} \neq 0] = P_x^{\pi}[X_t = 0, X_{t+1} = N_t]$; combining this equality with the inclusion

$$[X_t = 0, X_{t+1} = N_t] \subset [X_t = 0, X_{t+1} \neq 0]$$

it follows that the events $[X_t = 0, X_{t+1} \neq 0]$ and $[X_t = 0, X_{t+1} = N_t]$ differ by a null set with respect to P_x^{π} , and then

$$P_x^{\pi}[X_t = 0, X_{t+1} = N_t, \text{ i.o.}] = 1,$$

by Proposition 4.3(ii). Next, observe that the specification of the nonnegative reward function in (8), yields that

$$[X_t = 0, X_{t+1} = N_t] \subset \left[R(X_{t+1}, A_{t+1}) = N_t \right] \subset \left[\frac{\sum_{k=0}^{t+1} R(X_k, A_k)}{t+2} \ge \frac{N_t}{t+2} \right],$$

a relation that combined with the previous display leads to

$$P_x^{\pi} \left[\frac{\sum_{k=0}^{t+1} R(X_k, A_k)}{t+2} \ge \frac{N_t}{t+2} \text{ i.o.} \right] = 1,$$

and part (i) follows from this equality via Proposition 4.2.

(ii) Assume that $f_k \in \mathbb{F}$ is sample-path average optimal with optimal value g^* . In this case, the first part of Definition 3.1 and (15) together imply that $g^* \leq g$, and then the second condition in Definition 3.1 applied to the policy π in (18) yields that, for every $x \in S$,

$$\limsup_{n\to\infty}\frac{1}{n+1}\sum_{k=0}^n R(X_t,A_t) \le g^* \le g \quad P_x^{\pi}\text{-a.s.},$$

a statement that, since g < 1/2, contradicts part (i). Therefore, a sample-path average optimal stationary policy in the sense of Definition 3.1 does not exist.

5 Conclusion

This work considered discrete-time Markov decision chains on a denumerable state space. Besides standard continuity-compactness requirements, the main feature of the models analyzed in this note is that they admit a Lyapunov function. As already noted, in that context there exists a stationary policy which is optimal with respect to the expected average reward index; however, it was shown in Sect. 4 that Assumptions 2.1 and 2.2 do not imply the existence of a sample-path optimal stationary policy as specified in Definition 3.1, a conclusion that signals an interesting contrast between the expected and sample-path perspectives to the average criterion.

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